



ACHIEVING FAIRNESS WITH A SIMPLE RIDGE PENALTY ESTIMATES AND UNCERTAINTY

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December 12, 2025

→ INTRODUCTION

FAIR LINEAR MODELS

FAIR RIDGE REGRESSION

BAYESIAN FAIR REGRESSION

UNCERTAINTY QUANTIFICATION

CONCLUSIONS & ACKNOWLEDGEMENTS

- Machine learning (statistical?) models are being used in applications where it is crucial to ensure the **accountability and fairness** of the decisions made based on their outputs.
- Models are trained on historical data that contain various forms of bias, **capture those biases and carry them over** into current applications resulting in unfair discrimination of certain groups of people.
- The **concept of fairness** itself is difficult to define because it depends on the type of distortion we wish to limit and how we characterise it mathematically.
- How can we specify **fair models** that capture the non-discriminating information in the data and disregard the discriminating information?

Say that y is our response, $\hat{y}$ are fitted values from the model, S are the sensitive attributes containing the discriminating information and X are the other predictors.

- **Group fairness:** predictions should be similar across the groups identified by the sensitive attributes.
 - Statistical or demographic parity ($\hat{y} \perp\!\!\!\perp S$).
 - Equality of opportunity ($\hat{y} \perp\!\!\!\perp S \mid y$).
- **Individual fairness:** individuals that are similar receive similar predictions

$$f(y, S) = \sum_{i,j} d_1(y_i, y_j) d_2(s_i, s_j).$$

Many, many mathematical characterisations in the literature [10, 4, 11].

We can enforce fairness at different stages of the model selection, estimation and validation process [2]:

- **Pre-processing:** transform the data to remove the underlying discrimination so that models are guaranteed to be fair.
- **In-processing:** modify model estimation to remove discrimination, either by changing its objective function (typically the log-likelihood) or by imposing constraints on its parameters.
- **Post-processing:** assess a previously-estimated model, treated as a black box, and alter its predictions to make them fair.

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Komiyama *et al.* [7] did:

1. remove the association between $\mathbf{X}$ and $\mathbf{S}$ with $\mathbf{X} = \mathbf{B}^T \mathbf{S} + \mathbf{U}$, estimating $\widehat{\mathbf{B}}_{\text{OLS}} = (\mathbf{S}^T \mathbf{S})^{-1} \mathbf{S}^T \mathbf{X}$;
2. take the **decorrelated predictors** $\widehat{\mathbf{U}} = \mathbf{X} - \widehat{\mathbf{B}}_{\text{OLS}}^T \mathbf{S}$ which contain the component of $\mathbf{X}$ that cannot be explained by $\mathbf{S}$ ($\widehat{\mathbf{U}} \perp \mathbf{S}$);
3. formulate the regression model $\mathbf{y} = \mathbf{S}\boldsymbol{\alpha} + \widehat{\mathbf{U}}\boldsymbol{\beta} + \boldsymbol{\varepsilon}$;
4. formulate the fairness constraint

$$R_{\mathbf{S}}^2(\boldsymbol{\alpha}, \boldsymbol{\beta}) = \frac{\text{VAR}(\mathbf{S}\boldsymbol{\alpha})}{\text{VAR}(\widehat{\mathbf{y}})} = \frac{\boldsymbol{\alpha}^T \text{VAR}(\mathbf{S})\boldsymbol{\alpha}}{\boldsymbol{\alpha}^T \text{VAR}(\mathbf{S})\boldsymbol{\alpha} + \boldsymbol{\beta}^T \text{VAR}(\widehat{\mathbf{U}})\boldsymbol{\beta}};$$

5. solve the optimisation problem

$$\min_{\boldsymbol{\alpha}, \boldsymbol{\beta}} \mathbb{E}[(\mathbf{y} - \widehat{\mathbf{y}})^2] \quad \text{such that} \quad R_{\mathbf{S}}^2(\boldsymbol{\alpha}, \boldsymbol{\beta}) \leq r, r \in [0, 1].$$

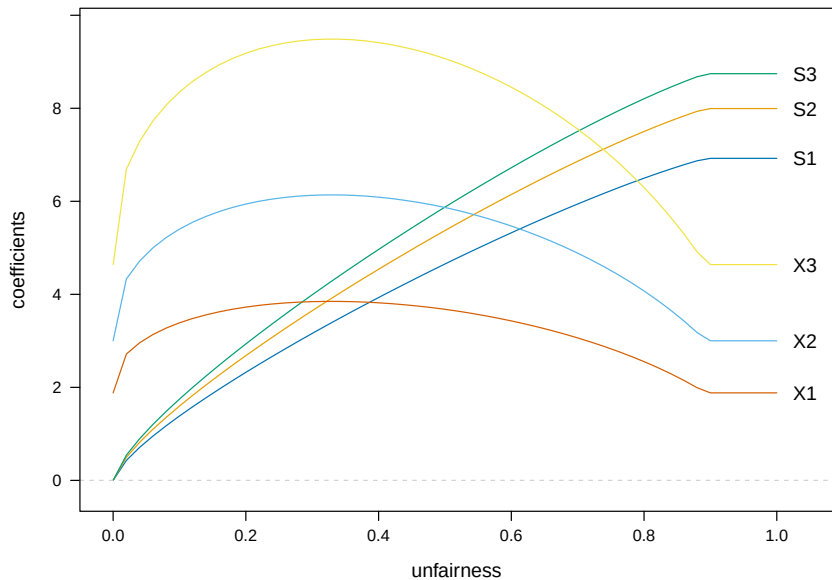
PROS:

- The formulation is simple.
- Discriminating and non-discriminating information are separated.
- The optimisation problem is QCQP, for which there are solvers.
- The fairness constraint is defined in terms of explained variance, the natural measure of information in a linear model.
- The bound is interpretable: 0 = complete fairness, 1 = no constraint.

CONS:

- No distributional assumptions.
- Cannot be extended without losing the ability to use QCQP solvers.
- The behaviour of the estimated coefficients is weird.

COEFFICIENT PROFILE PLOTS IN KOMIYAMA ET AL.



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Take two vintage pieces of statistics from the 1970s-1980s:

1. ridge regression (RR) [6];
2. generalised linear models (GLMs) [9].

With them, we build a **Fair (Generalised) Ridge Regression Model** (F(G)RRM) that fixes the CONS above and keep all the PROS:

- **Modular:** swappable characterisation of fairness.
- **Versatile:** supports all generalised linear models.
- **Interpretable:** both the model and the fairness constraints are interpretable, and all the best practices from the literature apply.
- **Statistical:** model selection, model validation, hypothesis testing, confidence intervals, etc. are already available in the literature.

Let's start again from $\mathbf{y} = \mathbf{S}\boldsymbol{\alpha} + \widehat{\mathbf{U}}\boldsymbol{\beta} + \varepsilon$. We want to re-create the shrinkage effects on the coefficients $\boldsymbol{\alpha}$ associated with $\mathbf{S}$ that we see in Komiyama *et al.*: we can do that with a ridge penalty,

$$(\widehat{\boldsymbol{\alpha}}_{\text{FRRM}}, \widehat{\boldsymbol{\beta}}_{\text{FRRM}}) = \underset{\boldsymbol{\alpha}, \boldsymbol{\beta}}{\operatorname{argmin}} \|\mathbf{y} - \mathbf{S}\boldsymbol{\alpha} - \widehat{\mathbf{U}}\boldsymbol{\beta}\|_2^2 + \lambda(r)\|\boldsymbol{\alpha}\|_2^2,$$

which we only apply to $\boldsymbol{\alpha}$ because by construction there is no discriminating information in $\widehat{\mathbf{U}}$. The parameter estimates are in closed form:

$$\begin{bmatrix} \widehat{\boldsymbol{\alpha}}_{\text{FRRM}} \\ \widehat{\boldsymbol{\beta}}_{\text{FRRM}} \end{bmatrix} = \begin{bmatrix} (\mathbf{S}^T\mathbf{S} + \lambda(r)\mathbf{I})^{-1} \mathbf{S}^T\mathbf{y} \\ (\widehat{\mathbf{U}}^T\widehat{\mathbf{U}})^{-1} \widehat{\mathbf{U}}^T\mathbf{y} \end{bmatrix}.$$

But how do we control the fairness of the model?

For a given level of fairness $r \in [0, 1]$:

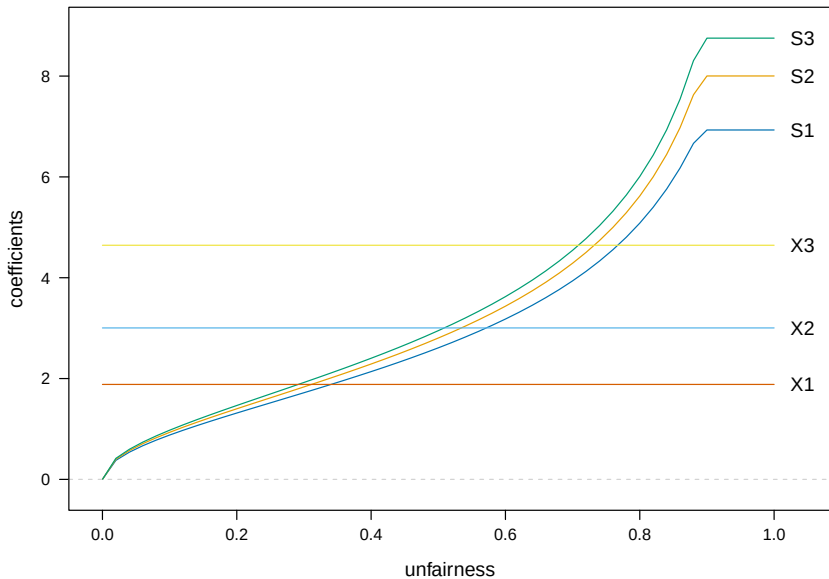
1. Compute $\widehat{\mathbf{U}}$ from $\mathbf{X}, \mathbf{S}$.
2. Estimate $\widehat{\boldsymbol{\beta}}_{\text{FRRM}} = (\widehat{\mathbf{U}}^T \widehat{\mathbf{U}})^{-1} \widehat{\mathbf{U}}^T \mathbf{y}$.
3. Estimate $\widehat{\boldsymbol{\alpha}}_{\text{OLS}} = (\mathbf{S}^T \mathbf{S})^{-1} \mathbf{S}^T \mathbf{y}$. Then:
 - 3.1 If $R_{\mathbf{S}}^2(\widehat{\boldsymbol{\alpha}}_{\text{OLS}}, \widehat{\boldsymbol{\beta}}_{\text{OLS}}) \leq r$, set $\widehat{\boldsymbol{\alpha}}_{\text{FRRM}} = \widehat{\boldsymbol{\alpha}}_{\text{OLS}}$.
 - 3.2 Otherwise, find the value of $\lambda(r)$ that satisfies

$$\boldsymbol{\alpha}^T \text{VAR}(\mathbf{S}) \boldsymbol{\alpha} = \frac{r}{1-r} \widehat{\boldsymbol{\beta}}_{\text{FRRM}}^T \text{VAR}(\widehat{\mathbf{U}}) \widehat{\boldsymbol{\beta}}_{\text{FRRM}}$$

and estimate the associated $\widehat{\boldsymbol{\alpha}}_{\text{FRRM}}$ in the process.

A single solution, requiring a simple **univariate** root finding algorithm regardless of the number of variables involved.

COEFFICIENTS IN FRRM



We can easily **replace $R_S^2(\alpha, \beta)$ with other constraints.**

1. $R_{EO}^2(\phi, \psi) = \frac{\text{VAR}(\mathbf{S}\phi)}{\text{VAR}(\mathbf{y}\psi + \mathbf{S}\phi)}$, from $\hat{\mathbf{y}} = \mathbf{y}\psi + \mathbf{S}\phi + \epsilon^*$.
 2. $f(\alpha, \mathbf{y}, \mathbf{S}) = \sum_{i,j} d(y_i, y_j)(\mathbf{s}_i\alpha - \mathbf{s}_j\alpha)^2$ and
 $D_{IF} = f(\hat{\alpha}_{\text{FRRM}}, \mathbf{y}, \mathbf{S}) / f(\hat{\alpha}_{\text{OLS}}, \mathbf{y}, \mathbf{S})$,
 3. Any convex combination of $R_S^2(\cdot)$, $R_{EO}^2(\cdot)$, $D_{IF}(\cdot)$ and others.
-

We can draw on [5, 12, 13] to estimate the **FGRRM**

$$(\hat{\alpha}_{\text{FRRM}}, \hat{\beta}_{\text{FRRM}}) = \underset{\alpha, \beta}{\operatorname{argmin}} D(\alpha, \beta) + \lambda(r) \|\alpha\|_2^2.$$

where $D(\cdot)$ is the **deviance** of a GLM + RR, choosing $\lambda(r)$ to give

$$\frac{D(\alpha, \beta) - D(\mathbf{0}, \beta)}{D(\alpha, \beta) - D(\mathbf{0}, \mathbf{0})} \leq r.$$

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A BAYESIAN CONSTRUCTION SIMILAR TO FGRRM?

Take a logistic FGRRM with linear component $\boldsymbol{\eta} = \mathbf{S}\boldsymbol{\alpha} + \widehat{\mathbf{U}}\boldsymbol{\beta}$:

$$\mathbf{y} \mid \widehat{\mathbf{U}}, \mathbf{S}, \boldsymbol{\alpha}, \boldsymbol{\beta} \sim \text{Ber}(\boldsymbol{\pi}), \quad \boldsymbol{\pi} = \text{logit}^{-1}(\boldsymbol{\eta}) = \frac{\exp(\boldsymbol{\eta})}{1 + \exp(\boldsymbol{\eta})}.$$

We can implement it with **Bayesian logistic regression** where

$$\boldsymbol{\alpha} \sim N\left(0, (\sigma_{\boldsymbol{\alpha}}^2/\lambda)\mathbf{I}_q\right), \quad \boldsymbol{\beta} \sim N\left(0, \sigma_{\boldsymbol{\beta}}^2\mathbf{I}_p\right),$$

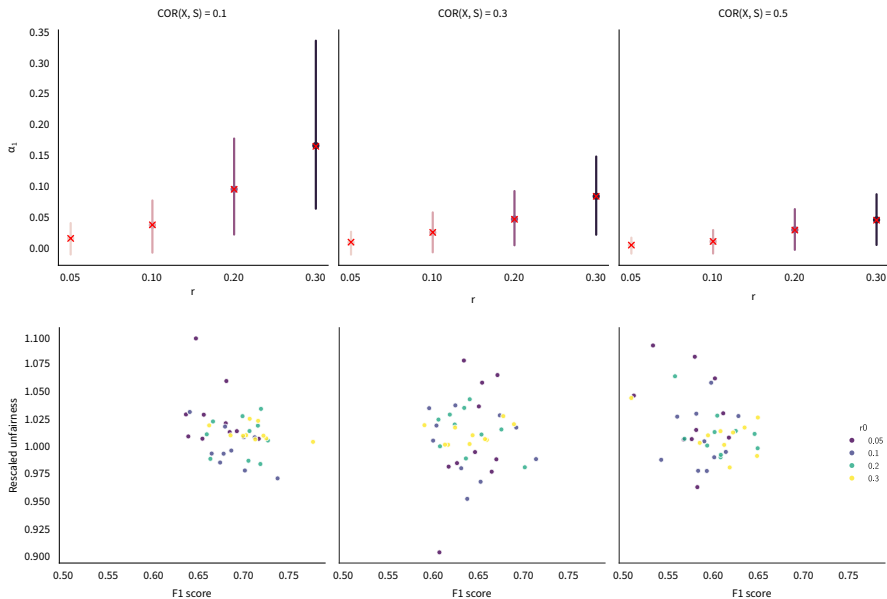
and tie everything together with a prior on $\lambda(r)$

$$P\left(\boldsymbol{\alpha}, \boldsymbol{\beta}, \lambda(r) \mid \mathbf{y}, \widehat{\mathbf{U}}, \mathbf{S}\right) \propto P\left(\mathbf{y} \mid \widehat{\mathbf{U}}, \mathbf{S}, \boldsymbol{\alpha}, \boldsymbol{\beta}\right) P\left(\boldsymbol{\beta}\right) P\left(\boldsymbol{\alpha} \mid \lambda(r)\right) P\left(\lambda(r)\right).$$

$\boldsymbol{\alpha}, \boldsymbol{\beta}$ and $\lambda(r)$ are stochastic, and induce a **posterior distribution on r** through the fairness constraint expression.

$P(r \mid \widehat{\mathbf{U}}, \mathbf{S} < r_0)$ then controls the fairness of the model.

MARGINAL PLOTS FROM THE POSTERIOR DISTRIBUTION



PROs:

- Putting Zellner's g-priors on α and β gives the MLE for $g \rightarrow \infty$, which is useful for comparisons.
- Posterior inference (jointly) on α , β and r via MCMC.
- Posterior sensitivity analysis on the choice of r .

CONs:

- It appears to be less numerically stable when $\mathbf{X}$ and $\mathbf{S}$ are collinear.
- MCMC convergence to its stationary distribution can be problematic in this and other less-than-smooth settings.

How can we approximate the MCMC inference in the original FGRRM?

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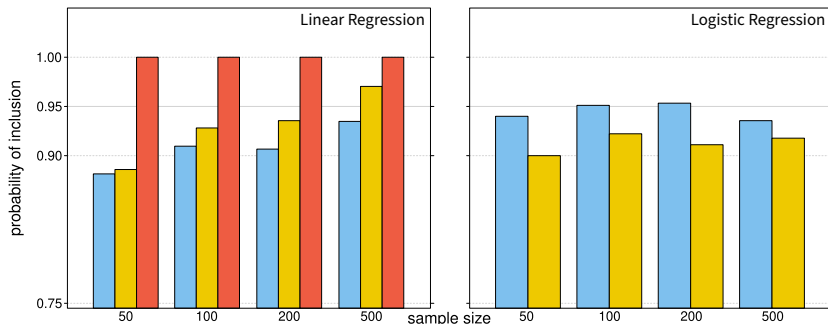
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WHAT CAN WE DO ON THE FREQUENTIST SIDE?

Much has been written on **confidence intervals for RR coefficients**.

For FGRRM:

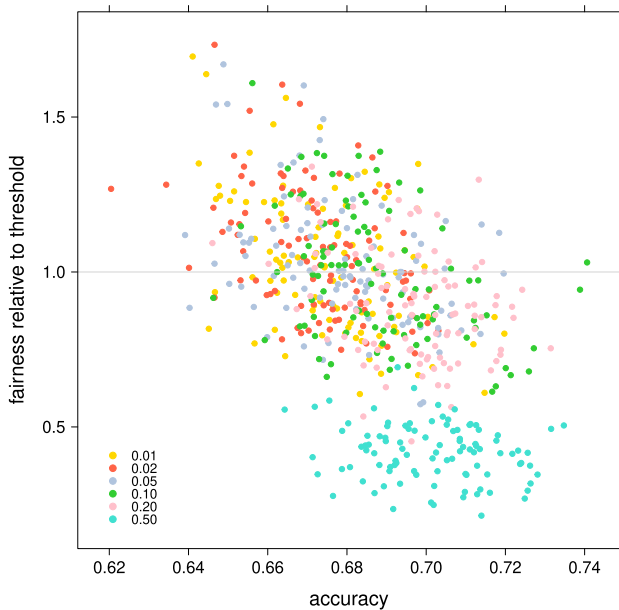
- Nonparametric bootstrap [3] works best.
- Double bootstrap [14] is costly, and works only for FRRM.
- Residual bootstrap [8] does not work.



At some computational cost, nested bootstrap allows us to **marginalise** $\alpha, \beta, \lambda(r)$ and training-validation split variability.

1. For $b = 1, \dots, B_1$:
 - 1.1 Draw a bootstrap sample $\mathbf{y}^{(b)}, \mathbf{X}^{(b)}$ and $\mathbf{S}^{(b)}$ from $\mathbf{y}, \mathbf{X}, \mathbf{S}$.
 - 1.2 Split training $(\mathbf{y}_{\text{TR}}^{(b)}, \mathbf{X}_{\text{TR}}^{(b)}, \mathbf{S}_{\text{TR}}^{(b)})$ and validation $(\mathbf{y}_{\text{VA}}^{(b)}, \mathbf{X}_{\text{VA}}^{(b)}, \mathbf{S}_{\text{VA}}^{(b)})$.
 - 1.3 For $b' = 1, \dots, B_2$:
 - 1.3.1 Draw a bootstrap sample $\mathbf{y}_{\text{TR}}^{(b')}, \mathbf{X}_{\text{TR}}^{(b')}, \mathbf{S}_{\text{TR}}^{(b')}$ from $\mathbf{y}_{\text{TR}}^{(b)}, \mathbf{X}_{\text{TR}}^{(b)}, \mathbf{S}_{\text{TR}}^{(b)}$.
 - 1.3.2 Estimate a model $\mathcal{M}^{(b')}$ with fairness r from $\mathbf{y}_{\text{TR}}^{(b')}, \mathbf{X}_{\text{TR}}^{(b')}, \mathbf{S}_{\text{TR}}^{(b')}$.
 - 1.3.3 Estimate the fairness loss for $\mathcal{M}^{(b')}$ on the validation set $(\mathbf{y}_{\text{VA}}^{(b)}, \mathbf{X}_{\text{VA}}^{(b)}, \mathbf{S}_{\text{VA}}^{(b)})$.
 - 1.3.4 Predict $\mathbf{y}_{\text{VA}}^{(b)}$ from $\mathbf{X}_{\text{VA}}^{(b)}, \mathbf{S}_{\text{VA}}^{(b)}$ using $\mathcal{M}^{(b')}$ to estimate predictive accuracy.
 2. Assess the joint distribution of empirical fairness and predictive accuracy either by visual inspection or using a pinball loss. [1]
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FAIRNESS-ACCURACY TRADE-OFF PLOT



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- Fairness is increasingly a concern as machine learning models become an integral part of automated decision support systems.
- Explainable AI investigates black-box models such as neural networks, but simpler models are also in common use and should be made fair.
- The literature studies fairness as an optimisation problem, producing models whose statistical properties and best practices are unknown.
- Classical statistics provides all the tools to formulate versatile fair models that we know how to interpret and use.
- Uncertainty quantification of fair models remains key, and is challenging because of the nature of the fairness constraint.



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THAT'S ALL!

HAPPY TO DISCUSS IN MORE DETAIL.

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