## What is the real prevalence of hypertension in France ? A hierarchical Bayesian modeling approach

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## Outline

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## Context

# Hypertension (HTN)

- Permanent high blood pressure (BP) level (if not treated)
  - $\rightarrow~$  Systolic/Diastolic BP  $\geq 140/90~mmHg$
- Leading modifiable risk factor for cardiovascular and renal diseases
- Most frequent chronic disease  $\rightarrow$  majour issue in terms of resources allocation

## Diagnosis of HTN

- Clinical diagnosis based on multiple BP measurements during several visits
  - ightarrow Control for within subject variability
- In epidemiological studies, BP usually measured during a single visit (cost++)
  - $\Rightarrow\,$  Biased estimates of HTN if within-person variability neglected ^1
    - $\rightarrow~$  Correction made using within-person variability estimates from external studies
    - $\rightarrow\,$  Correction depends on the composition of the population of external studies (i.e. age and sex)

## Objectives

- 1. Propose a method of correction that takes into account the main factors influencing BP variability : age and sex
- 2. Apply the method to estimate HTN prevalence in France in 2015

1. O. H. KLUNGEL et al. "Estimating the prevalence of hypertension corrected for the effect of within-person variability in blood pressure". eng. Journal of Clinical Epidemiology 53.11 (nov. 2000), p. 1158-1163.

#### Notations and distributional assumptions

Components of BP measures

#### For a given sex and type of BP

Let  $y_{ivm}$  denote the  $m^{\text{th}}$  measure of blood pressure for the patient *i* of age  $a_i = a$ , during the visit *v*.

$$y_{ivm} = f(a) + u_i + v_{iv} + \epsilon_{ivm} \tag{1}$$

where

- f(a) : mean BP level for population of age a

Indivudal BP level 
$$y_i = f(a) + u_i$$

- $u_i$ : deviation from f(a) for individual i
- $v_{iv}$  : deviation from individual BP level during visit v
- $\epsilon_{ivm}$  : measurement error of the  $m^{\text{th}}$  measure during the visit v

 $u_i, v_{iv}$  and  $\epsilon_{ivm}$  considered as iid gaussian random fluctuations, with variances depending on age :

$$\left\{egin{aligned} u_i &\sim \mathcal{N}(0, g(a)) \ v_{i v} &\sim \mathcal{N}(0, h(a)) \ \epsilon_{i v m} &\sim \mathcal{N}(0, l(a)) \end{aligned}
ight.$$

These assumptions imply a normal distribution for  $y_{ivm}$ .

htn = proportion of individuals with individual BP level  $y_i$  above a threshold

 $(y_i|a_i = a) = f(a) + u_i \sim \mathcal{N}(f(a), g(a))$ 



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#### Natural estimator for $y_i$

For v visits and m measures per visit :

$$- (\bar{y}_i | a_i = a) = f(a) + u_i + \frac{1}{v} \sum_{k=1}^{v} v_{ik} + \frac{1}{mv} \sum_{k=1}^{v} \sum_{l=1}^{m} \epsilon_{ikl}$$

$$- V(\bar{y}_i|a_i = a) = g(a) + \frac{1}{v}h(a) + \frac{1}{mv}I(a) > V(y_i|a_i = a)$$

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- $-V(\bar{y}_i|a_i = a) = g(a) + \frac{1}{v}h(a) + \frac{1}{mv}I(a) > V(y_i|a_i = a)$

#### Plug $\bar{y}_i$ in place of $y_i$ in (2) is biased

- Direction of the bias depends on the sign of T f(a) (i.e. positive bias if T > f(a), negative otherwise)
- Magnitude increases with  $\frac{1}{v}h(a) + \frac{1}{mv}l(a)$

Corrected estimator

Rescale  $\bar{y}_i$  so that the resultant has the expected variance Defining c(a): correction factor for age a

$$ightarrow c(a) = \sqrt{rac{g(a)}{V(ar{y}_i)}} = \sqrt{rac{g(a)}{g(a) + rac{1}{v}h(a) + rac{1}{mv}l(a)}}$$

Then

$$y_i^c = f(a) + c(a)(\bar{y}_i - f(a))$$
 (3)

has a gaussian distribution with mean f(a) and variance g(a).

 $\rightarrow \hat{y}_i^c : y_i^c$  estimated by substituting f(a) by  $\frac{1}{n(a)} \sum_{a_i=a} \bar{y}_i$  in (3).

#### Corrected estimator

$$\hat{htn}(a) = rac{1}{n(a)} \sum_{a_i = a} \mathbb{1}_{\hat{y}_i^c > T}$$

But we don't know c(a)...

## Klungel - I

Use correction factor from other studies!

Study	N	Age	% women	Visits	Measure	Time v	Time m
Klungel	834	20–59	50%	2	2	1 y	5 min
Rosner 1	991	30-69	47%	3/4/5	3	1/7  days	30 sec
Rosner 2	326	0–69	37%	2/3	3	1 week	30 sec
Hebel	100	30-69	50%	2	2	3 years	5 min
Cook	2,111	16 - 49	100%	2	3	3 years	30 sec
Hughe	11,299	30-59	0%	4	1	1 year	-
Armitage	50	47.6	0%	4	1	1 year	-



## Klungel - II

- Use a single mean correction factor (lack of detailed data)
- Corrections factors vary according to
  - The delay between visits
  - The number of measurement within visit
  - The age and sex composition of the studied population

#### Room for some improvement

- $\rightarrow\,$  Derive general shapes of the components of BP variability, by age and sex
- $\rightarrow\,$  Correction factor by age and sex

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## Estimation of c(a) I

Estimation of the components of variance of yivm

$$c(a) = \sqrt{\frac{g(a)}{g(a) + \frac{1}{v}h(a) + \frac{1}{mv}l(a)}}$$

#### What is needed to estimate c(a)

Components of variability of  $y_{ivm}$ :

- g(a) : variability of  $y_i$  across individuals
- h(a): variability of BP between visits within an individual
- I(a) : variability of the measures of BP within an individual during the same visit

#### Need data with multiple measure of BP during several visits

#### How to estimate the components

- ANOVA like estimates
- Hierarchical Bayesian linear models

## Estimation of c(a) II

Estimation of the components of variance of yivm

## Hierarchical model $y_{ivm} = f(a) + u_i + v_{iv} + \epsilon_{ivm}$ with $u_i \sim \mathcal{N}(0, g(a)), v_{iv} \sim \mathcal{N}(0, h(a))$ and $\epsilon_{ivm} \sim \mathcal{N}(0, l(a))$

- Specification of random effects (example of  $u_i$ ) :
  - ightarrow Random intercept by individual  $u^s_i \sim \mathcal{N}(0,1)$
  - ightarrow Multiplied by a positive scale parameter depending on age :  $\exp(g^s(a))$

$$\Rightarrow u_i = u_i^s \exp(g^s(a)) \Rightarrow V(u_i) = [\exp(g^s(a))]^2 = g(a)$$

- Same for  $v_{iv} = v^s_{iv} \exp(h^s(a))$  and  $\epsilon_{ivm} = \epsilon^s_{ivm} \exp(l^s(a))$
- f(a),  $g^{s}(a)$ ,  $h^{s}(a)$ , and  $l^{s}(a)$  estimated with penalized thin plate splines<sup>2</sup>:

For a function k(a) :  $k(a) = \alpha + \beta a + \sum_j b_j z_j(a)$ 

- $\rightarrow z_j(a)$  : known splines basis function
- ightarrow ~eta and  $b_j$  parameters to be estimated

 $\rightarrow$  Penalization of wiggliness by imposing a gaussian prior on the  $b_j$  :  $b_j \stackrel{iid}{\sim} \mathcal{N}(0, au)$ 

<sup>2.</sup> Simon N. WOOD. "Stable and Efficient Multiple Smoothing Parameter Estimation for Generalized Additive Models". Journal of the American Statistical Association 99.467 (sept. 2004), p. 673-686.

# Estimation of c(a) III

Estimation of the components of variance of yivm

#### Priors

Following Gelman's<sup>3</sup> recommendations (default brms priors<sup>4</sup>)

- Intercepts in  $g^{s}(a)$ ,  $h^{s}(a)$ , and  $l^{s}(a)$ : centered Student distribution with 3 degree of freedom and a scale of 2.5
- Intercepts in f(a) :  $\mathcal{N}(0, 10000)$
- Linear fixed effects : improper flat prior over the reals
- Standard deviations (i.e. penalties for splines) : half student-t prior with 3 degrees of freedom and a scale of 2.5.

#### Estimation

- Hamiltonian Monte Carlo with Stan software<sup>5</sup>
- 4 chains with 6000 iteration (5000 burn-in)
- Numerical computations performed on the S-CAPAD/DANTE platform, IPGP, France

<sup>3.</sup> Andrew GELMAN. "Prior distributions for variance parameters in hierarchical models (comment on article by Browne and Draper)". *Bayesian Analysis* 1.3 (sept. 2006), p. 515-534.

<sup>4.</sup> Paul-Christian BÜRKNER. "brms : An R Package for Bayesian Multilevel Models Using Stan". Journal of Statistical Software 80.1 (2017), p. 1-28.

<sup>5.</sup> STAN DEVELOPMENT TEAM. stan Modeling Language Users Guide and Reference Manual, 2.27. 2021.

#### Estimation of c(a) IV Data

#### Data from NHANESIII study - 1988-1994

- 18,825 adults from general US population (  $\geq$  17 y.o.)
- 2 or 3 (n=2,174) visits, 2 BP measurements per visit
- First visit to a mobile examination center
- Median duration between subsequent consecutive of 17 days (minimum 1 day, maximum 48 days)

Hierarchical model estimated separately by :

- Sex
- Type of blood pressure (i.e. systolic and diastolic)
- Patients taking or not anti-hypertensive treatments

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#### Results - convergence

Traceplots of model parameters for systolic blood pressure in men - untreated patients



# Results - convergence $\hat{R}$ and ESS - untreated patients

			Men		Women			
BP type	gp	par	Ŕ	ess (bulk)	ess (tail)	Ŕ	ess (bulk)	ess (tail
Diastolic	f(a)	α	1.00	1,399.30	2,160.29	1.00	2,121.48	2,380.68
	. ,	β	1.00	1,304.41	1,681.94	1.00	2,379.13	2,917.31
		$\sqrt{\tau}$	1.00	1,014.70	1,407.12	1.00	1,575.97	2,183.74
	l(a)	à	1.00	2,214.12	2,910.38	1.00	1,939.27	2,856.58
		β	1.00	1,563.06	2,166.84	1.00	2,491.48	2,484.32
		$\sqrt{\tau}$	1.00	1,214.76	2,023.57	1.00	1,133.40	2,119.20
	g(a)	$\dot{\alpha}$	1.01	662.16	1,325.92	1.00	955.40	1,959.27
		β	1.00	980.40	922.60	1.00	1,237.86	1,822.75
		$\sqrt{\tau}$	1.00	660.84	1,196.00	1.00	846.64	1,335.66
	h(a)	$\dot{\alpha}$	1.00	940.02	1,716.07	1.00	1,096.93	1,900.19
		β	1.00	1,317.91	1,901.86	1.00	748.18	1,513.1
		$\sqrt{\tau}$	1.00	927.64	1,808.28	1.01	673.37	1,029.3
Systolic	f(a)	α	1.00	1,045.55	1,952.33	1.00	2,082.74	2,349.8
		$\beta$	1.00	1,345.91	1,894.80	1.00	1,997.69	2,079.7
		$\sqrt{\tau}$	1.01	1,107.39	2,051.83	1.00	1,786.75	2,272.82
	l(a)	$\alpha$	1.00	2,183.92	3,091.98	1.00	1,357.95	2,692.41
		$\beta$	1.00	2,220.05	2,775.54	1.00	2,684.52	2,776.89
		$\sqrt{\tau}$	1.00	1,679.53	2,336.00	1.00	2,211.08	2,568.53
	g(a)	$\alpha$	1.00	1,084.71	1,921.31	1.00	1,596.61	2,204.35
		$\beta$	1.01	899.73	1,694.29	1.00	1,823.53	2,636.01
		$\sqrt{\tau}$	1.01	600.28	1,267.82	1.00	1,907.28	2,238.26
	h(a)	$\alpha$	1.00	1,072.68	2,223.07	1.00	1,206.96	2,391.47
		$\beta$	1.00	1,171.63	1,426.56	1.00	1,958.17	2,766.99
		$\sqrt{\tau}$	1.00	908.50	1,736.86	1.00	2,094.04	2,719.96

# Results (untreated patients)

Components of variance : Systolic blood pressure



# Results (untreated patients)

Components of variance : Diastolic blood pressure\_



#### HTN prevalence in France Application to ESTEBAN data

#### ESTEBAN study

- Cross-sectional study (2014-2016)
- $\sim$  2,000 individuals 18 to 74 y.o.
- 2 BP measures during a single visit

#### Estimation

- 1. For each post-sample of c(a)
  - Correct individual BP  $\rightarrow y_i^c$
  - Estimate HTN (using sampling weights)
    - $y_i^c > \text{threshold OR}$
    - Patient treated for HTN
- 2. Combine post-sample's estimates
  - Variance = mean variance of post-samples + variance across post-sample's estimates

## HTN prevalence in France

Prevalence by age



#### Overall

	Un-corrected	Corrected
Men Women	38.1[34.2 ;42.2] 25.0[21.9 ;28.4]	35.0[31.2;39.1] 21.3[18.4;24.5]
All	31.3[28.8;34.0]	27.9[25.5;30.5]

- Larger differences in women than in men
- Larger differences in young than in elderly
- $\rightarrow~$  12.7 instead of 14.3 millions of cases for the 18-74

## HTN prevalence in France

Effect of correction in sub-pops



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#### Discussion

#### Method

- Control for differences in age and sex composition of study (e.g. ESTEBAN) vs reference study (e.g. NHANESIII)
  - Main factors driving variability of BP
  - Easy to apply to subpop
- Main hypothesis : c(a) estimated from external data applies to study
  - $\rightarrow~$  Less restrictive than equality of variances
  - $\rightarrow\,$  Compatibility between populations/study protocol ?
  - $\rightarrow$  In our case, the variability of  $y_{ivm}$  observed in ESTEBAN  $\simeq$  predicted from NHANESIII component of variance
- Other factors influencing BP variability not accounted for

## Hierarchical modeling

- Gaussian assumption
- No correlation between components of variance

#### Discussion

#### Results

- Substantial variations of c(a) with age and sex
- Modest to substantial correction of HTN
- Within CI bands

## R package

Method of correction disseminated in a R package available at https://github.com/echatignoux/BPpack

Appendix



## Mean BP levels in Esteban vs Mean BP levels predicted from NHANES

## Total variance in Esteban vs total variance predicted from NHANES



## Measurement error in Esteban vs measurement error predicted from NHANES



#### ANOVA derivation of variance components

If m measures of BP are realized during v visits, analytical estimator of h, g and l can be derived using an ANOVA approach.

If we note  $\bar{y}_{iv}$  the mean of BP measures for individual *i* during visit *v*, then  $V(y_{ivm} - \bar{y}_{iv}) = \frac{m-1}{m}I(a)$ , leading to

$$\hat{l}(a) = rac{m}{m-1}V(y_{ivm}-ar{y}_{iv}|a_i=a)$$

Similarly,  $V(y_{ivm} - \bar{y}_i) = \frac{v-1}{v}h(a) + \frac{mv-1}{mv}I(a)$ , so h(a) may be estimated by

$$\hat{h}(\mathsf{a}) = rac{v}{v-1} V(y_{ivm} - ar{y}_{iv}|\mathsf{a}_i = \mathsf{a}) - rac{m-1}{m(v-1)} \hat{l}(\mathsf{a})$$

An estimator for g(a) derives from the expressions above :

$$\hat{g}(a) = V(y_{ivm}|a_i = a) - \hat{h}(a) - \hat{l}(a)$$