
L'astuce bayésienne en test multiple

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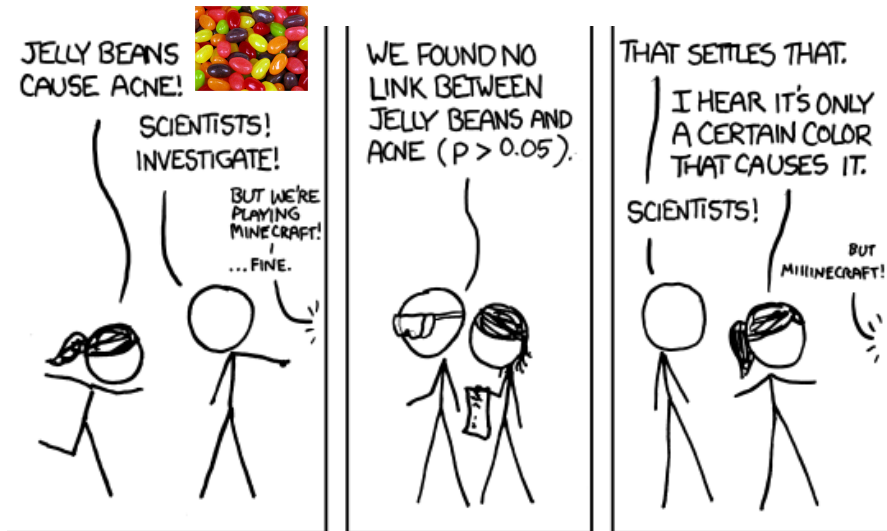
Journée AppliBUGS 2016

1 Introduction

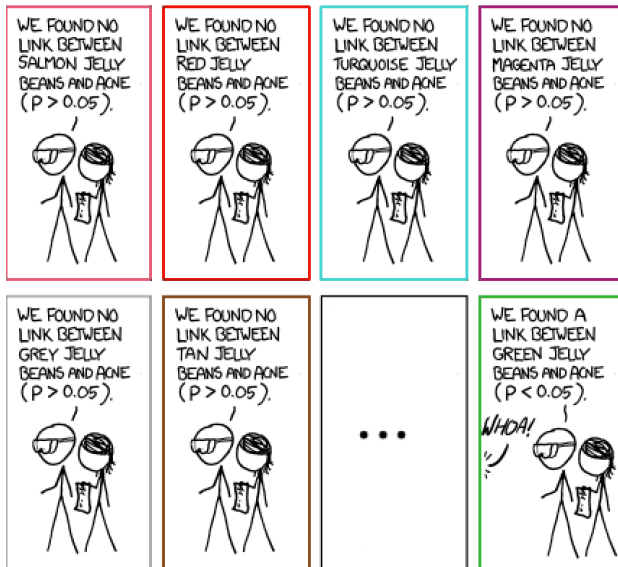
2 BH procedure

3 Bayesian tricks

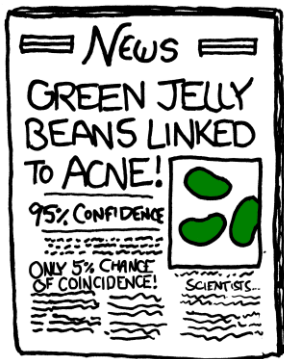
A "multiple testing story" (<http://xkcd.com>)



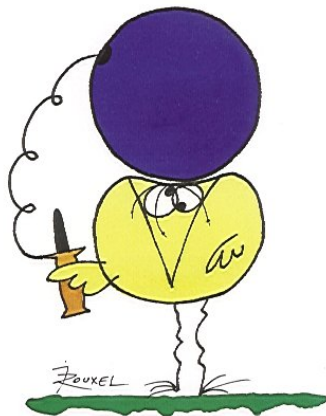
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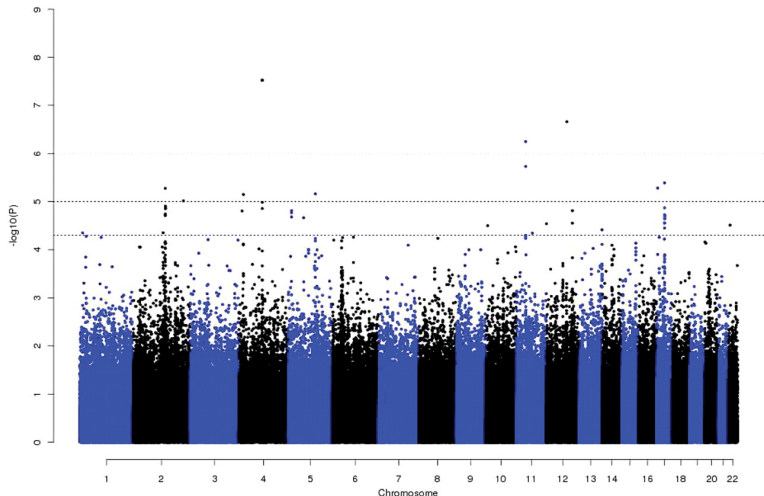


Les devises Shadok



EN ESSAYANT CONTINUELLEMENT
ON FINIT PAR RÉUSSIR. DONC:
PLUS ÇA RATE, PLUS ON A
DE CHANCES QUE ÇA MARCHE.

Manhattan plot (GWAS data)



- ▶ Find interesting SNPs
- ▶ Appropriate threshold ?

Simple multiple testing setting

Model:

- ▶ $\theta = (\theta_1, \dots, \theta_m) \in \{0, 1\}^m$ arbitrary
- ▶ $X = (X_1, \dots, X_m) \mid \theta$ mutually independent,

$$X_i \sim \mathcal{N}(0, 1) \text{ if } \theta_i = 0$$

$$X_i \sim \mathcal{N}(\mu, 1) \text{ if } \theta_i = 1,$$

for some $\mu > 0$.

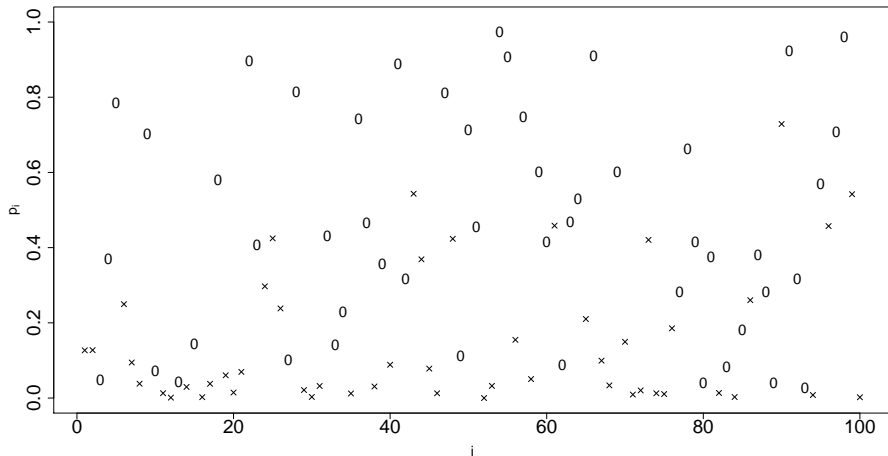
Aim: recover θ from $(X_i, 1 \leq i \leq m)$

Normalization:

- ▶ p -values

$$p_i(X) = \bar{\Phi}(X_i), \quad 1 \leq i \leq m$$

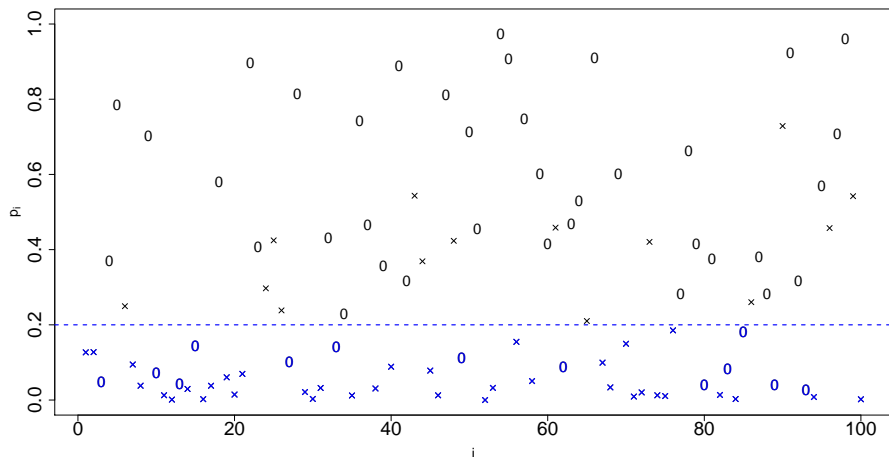
Thresholding



▶ Rejection set $R = \{1 \leq i \leq m : p_i(X) \leq t\}$

▶ Some false positives

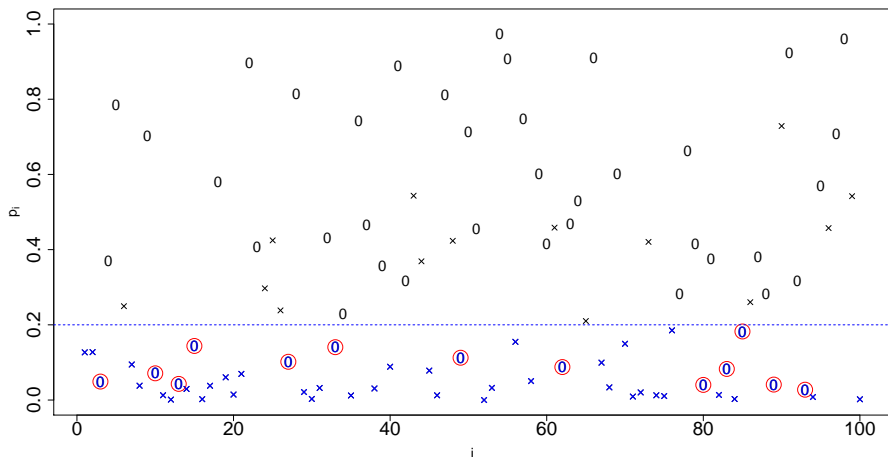
Thresholding



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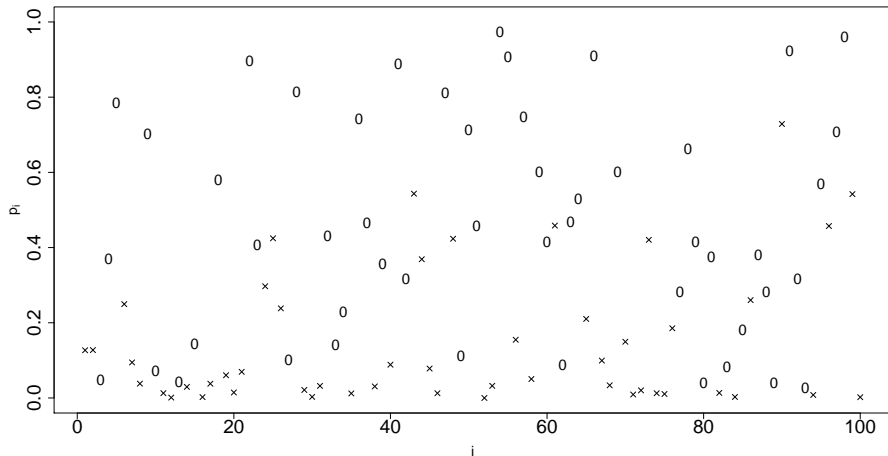
► Some false positives

Thresholding



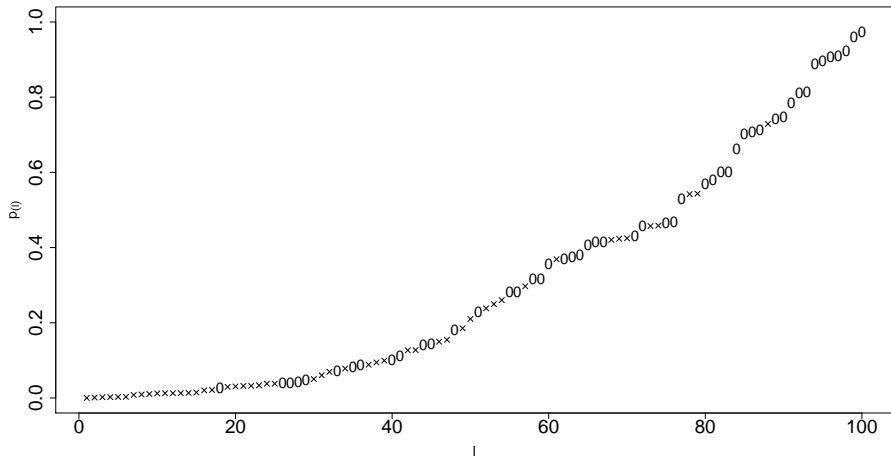
- ▶ Rejection set $R = \{1 \leq i \leq m : p_i(X) \leq t\}$
- ▶ Some **false positives**

Ordering p -values



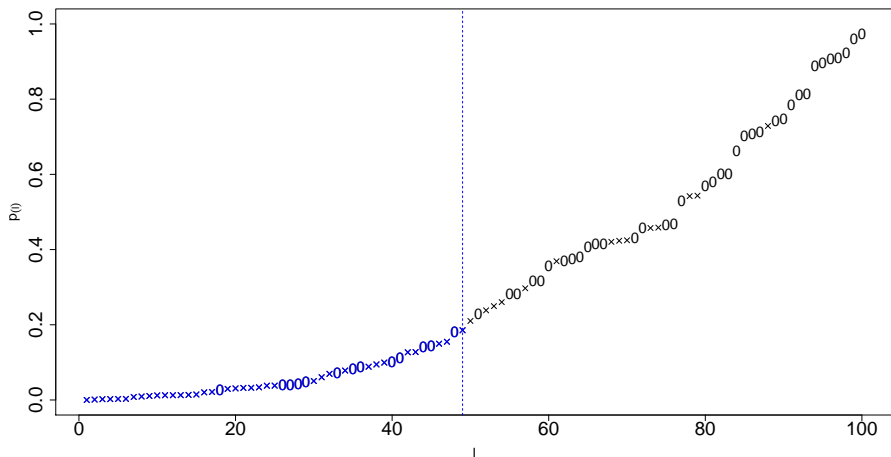
- ▶ Stopping rule
- ▶ Reject $\hat{\ell}$ smallest p -values

Ordering p -values



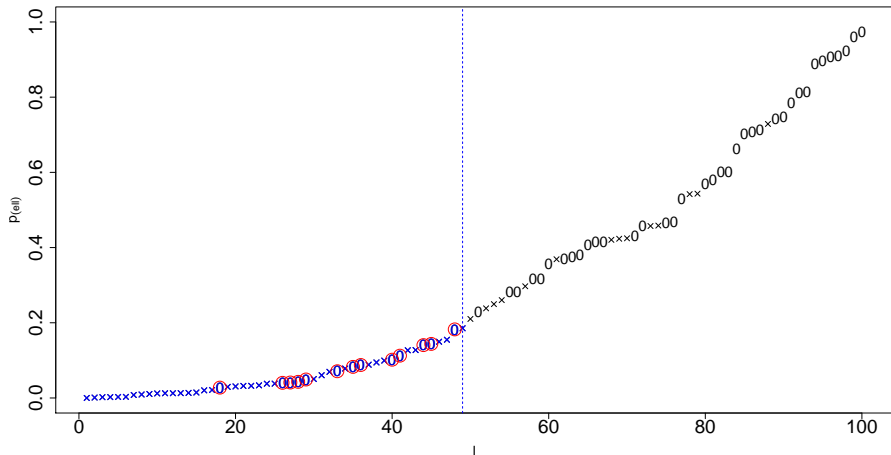
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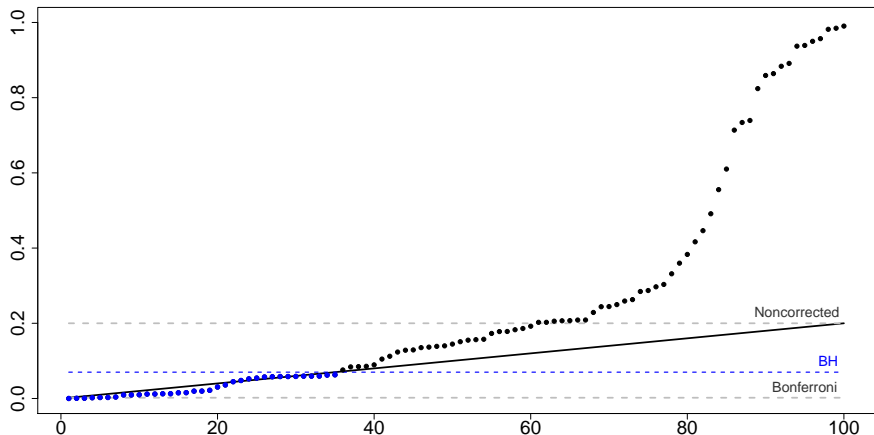
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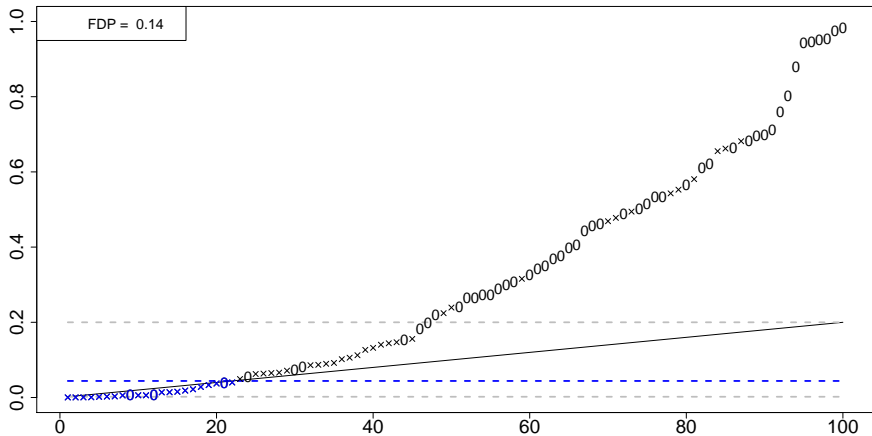
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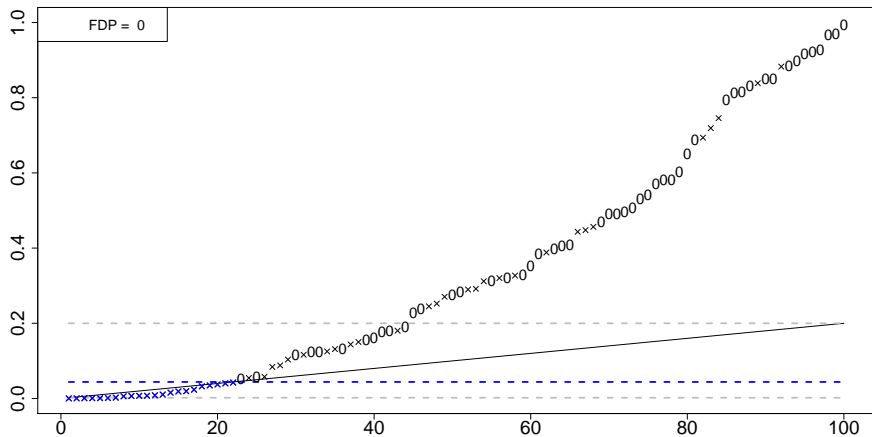
$$\hat{t} = \alpha(\hat{k} \vee 1)/m \quad \text{with} \quad \hat{k} = \max\{0 \leq k \leq m : p_{(k)} \leq \alpha k/m\}$$



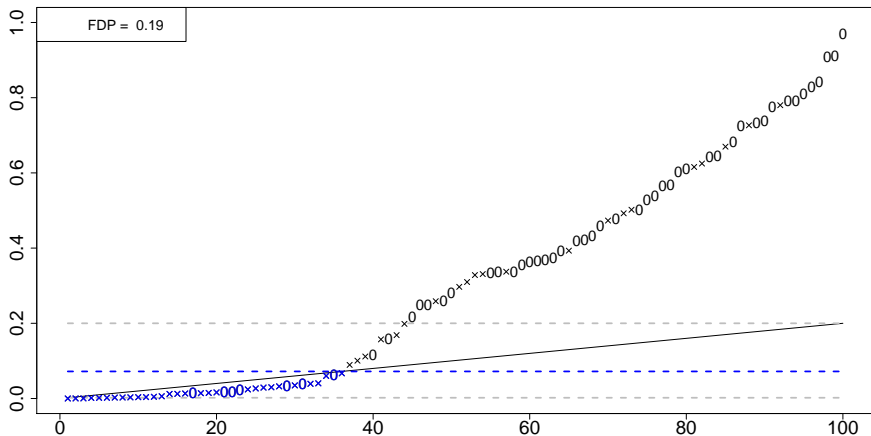
Illustration, $\alpha = 0.2$, $m = 100$



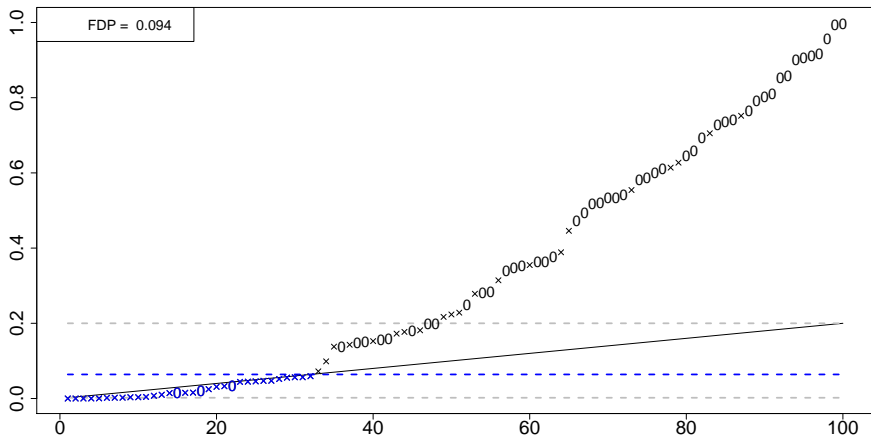
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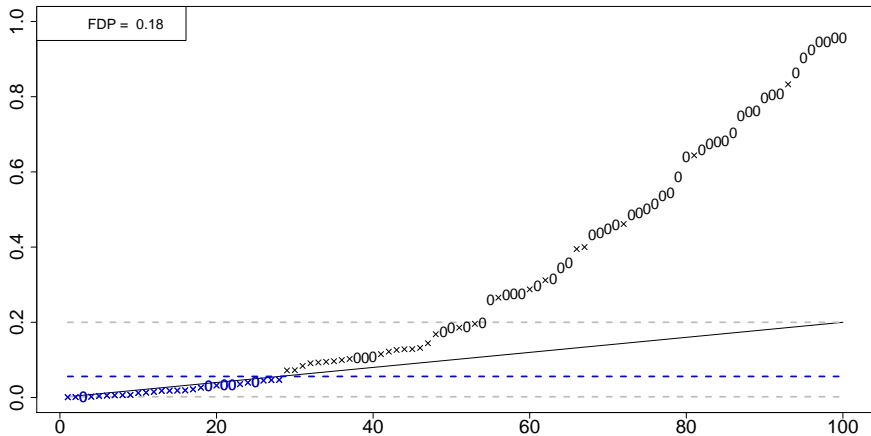
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Illustration, $\alpha = 0.2$, $m = 100$



More formally ...

False discovery rate

For a threshold t , $V(t) = \sum_{i=1}^m (1 - \theta_i) \mathbf{1}\{p_i \leq t\}$, $R(t) = \sum_{i=1}^m \mathbf{1}\{p_i \leq t\}$,

$$\text{FDR}(t, P) = \mathbf{E}[\text{FDP}(t, P)], \quad \text{FDP}(t, P) = \frac{V(t)}{R(t)} \quad \left(\frac{0}{0} = 0 \right)$$

here $P = (\theta, \mu)$ parameter.

Difficult point: random denominator

Theorem [Benjamini and Hochberg (1995)]

$$\text{for all } \theta, \mu, \quad \text{FDR}(\text{BH}_\alpha, (\theta, \mu)) \leq \alpha$$

Widely used

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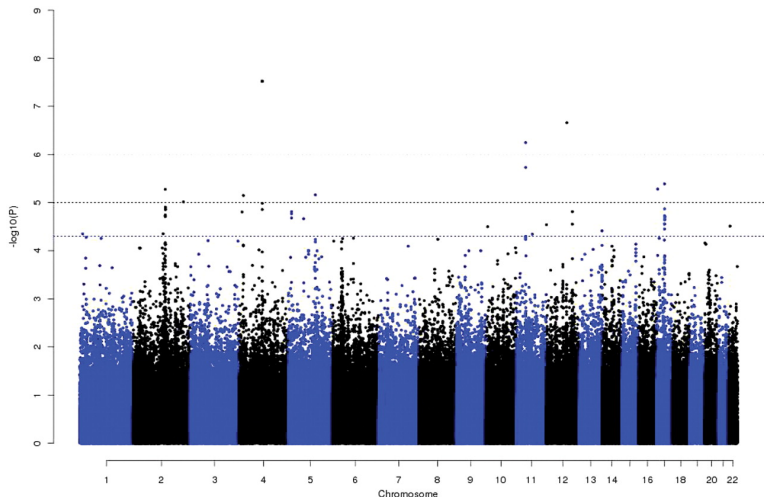
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Manhattan plot (GWAS data)



- ▶ Looking for **structure** on $\theta_1, \dots, \theta_m$
- ▶ Structure = a priori

Bayesian multiple testing literature

[Efron and Tibshirani. *Empirical Bayes methods and false discovery rates for microarrays* (2002, Gen. Epi.)]

- ▶ a priori : $\theta_1, \dots, \theta_m$ i.i.d. (unstructured)
- ▶ BH 'optimal' (empirical Bayes method)

[Sun and Cai. *Large-scale multiple testing under dependence* (2009, JRSSB)]

- ▶ $\theta_1, \dots, \theta_m$ with a Markov structure
- ▶ New Bayesian multiple testing procedures
- ▶ BH suboptimal

(See also [Storey (2003)], [van der Laan et al. (2005)], [Scott and Berger (2005)], [Tang et al. (2007)], [Ghosal and Roy (2012)] ...)

First trick: optimal classification

Model:

- ▶ $\theta = (\theta_1, \dots, \theta_m) \in \{0, 1\}^m$ Markov chain transition matrix A
- ▶ $X = (X_1, \dots, X_m) \mid \theta$ mutually independent,

$$X_i \sim \mathcal{N}(0, 1) \mid \theta_i = 0$$

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- ▶ p -values

$$p_i(X) = \bar{\Phi}(X_i), \quad 1 \leq i \leq m$$

- ▶ ℓ -values

$$\ell_i(X) = \mathbf{P}(\theta_i = 0 \mid X), \quad 1 \leq i \leq m$$

Theorem:

mis-classification risk minimum for the rule

$$\Phi_i(X) = \mathbf{1}\{\ell_i(X) \leq 1/2\}, \quad 1 \leq i \leq m$$

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▶ $m = 200$

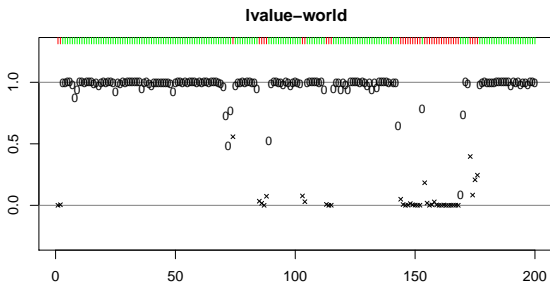
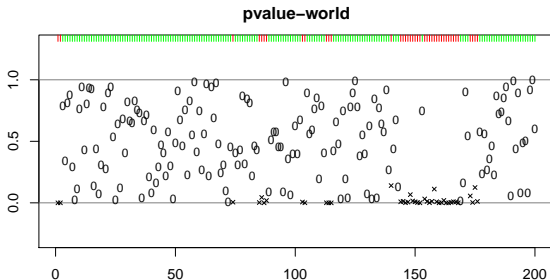
▶ Transition matrix

$$A = \begin{pmatrix} 0.95 & 0.05 \\ 0.25 & 0.75 \end{pmatrix}$$

▶ Gaussian

$$\mu = 2.5$$

$$\sigma = 1$$

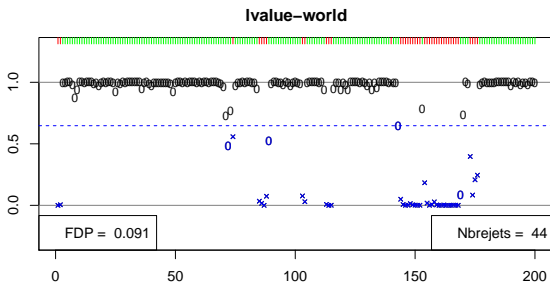
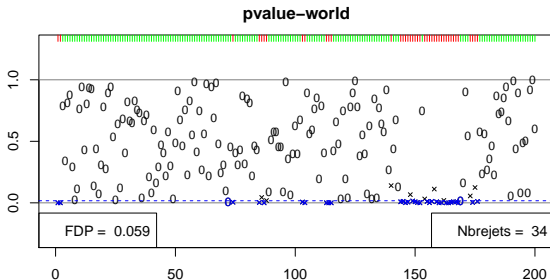


Second trick: Bayesian FDR

▶ $\alpha = 0.1$

▶ BH procedure

▶ Bayesian procedure



Second trick: Bayesian FDR

Bproc : reject \hat{k} smallest ℓ -values for

$$\hat{k} = \max \left\{ k : \frac{1}{k} \sum_{k'=1}^k \ell_{(k')} \leq \alpha \right\}.$$

Theorem:

$$\text{for all } \mu, \quad \int_{\theta} \text{FDR}(\text{Bproc}, (\theta, \mu)) d\nu(\theta) \leq \alpha$$

Proof : "one line"

Limitations

- ▶ FDR control is only valid under the prior
- ▶ ℓ -values to be computed
- ▶ procedure depends on hyper-parameters

If a priori assumed 'true' :

- ▶ Hyper-parameters have to be estimated (empirical Bayes)
- ▶ Plug-in step needs additional theory

If a priori just a tool:

frequentist FDR control of Bayesian multiple testing procedures?

for all $\theta \in \Theta, \mu, \text{FDR}(\text{Bproc}, (\theta, \mu)) \approx \alpha$ under sparsity

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