

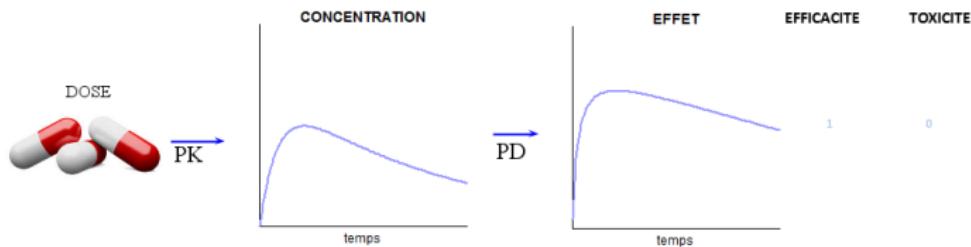
Bayesian approaches for pharmacogenetic models with JAGS and Stan

Julie Bertrand, Maria De Iorio, David J. Balding



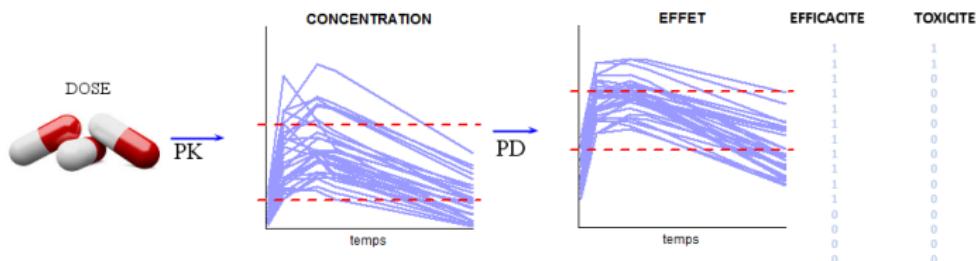
Pharmacological and genetic variability

- Clinical pharmacology: study the interaction between the organism and the drug
 - pharmacokinetics (PK) and pharmacodynamics (PD)



Pharmacological and genetic variability

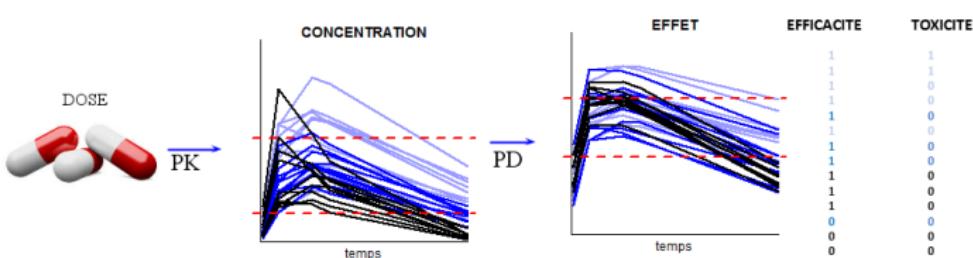
- Clinical pharmacology: study the interaction between the organism and the drug and its variability
 - pharmacokinetics (PK) and pharmacodynamics (PD)



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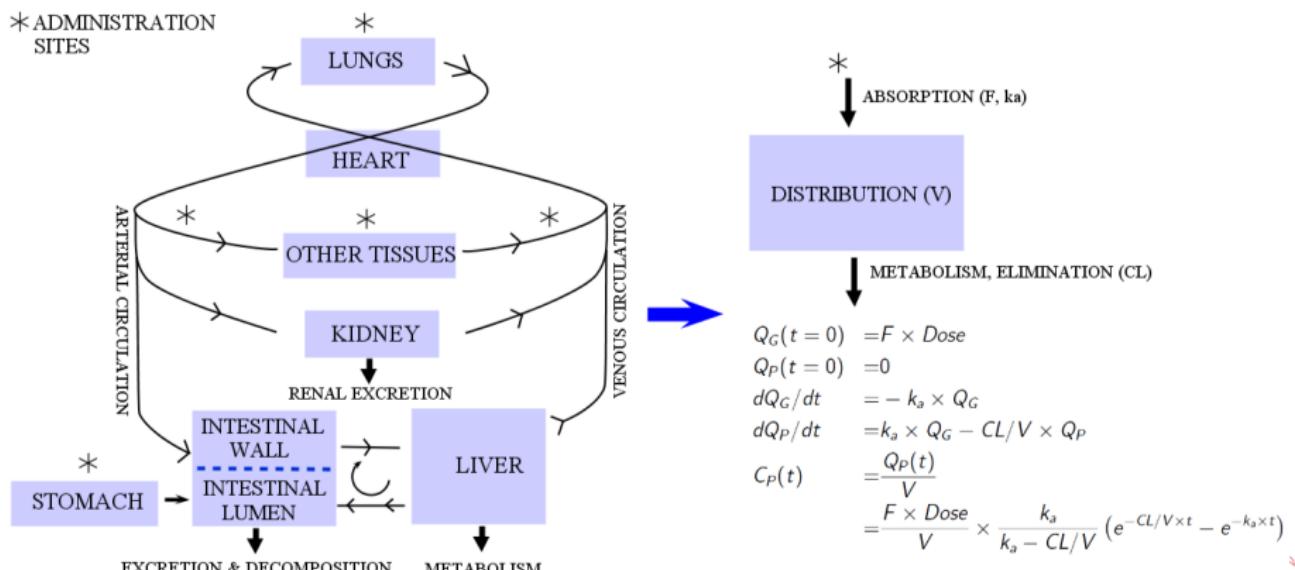
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TT	TT	CC	AA	AA	TT	GG	TT	TT	TT
TT	TT	TT	GG	GG	CC	GG	GG	GG	GG
CT	TT	CT	AG	AG	CT	GG	GT	GT	GT
CT	TT	CC	AA	AA	TT	GG	GT	GT	GT
CT	TT	TT	GG	GG	CC	GG	GG	GG	GG



- Pharmacogenetics (PG): genetic part of the variability
 - stratified medicine
- Genes coding for proteins involved in PK/PD processes
 - metabolism enzymes (CYP450, NAT)
 - single nucleotide polymorphism (SNP)

Modelling in pharmacology

- Semi-physiological models integrating the *a priori* knowledge on the drug
 - parameters characterizing each physiological processes
 - model nonlinear in its parameters



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Example : $CL_i = CL + \beta \times SNP_i + \eta_i$ with $SNP = \{0, 1, 2\}$

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- Estimation methods
 - Maximum likelihood: model linearization, Gaussian quadrature, SAEM
 - Bayesian inference (Lunn et al. 2002)

Methodological challenges in PGx

- PK/PD phenotype → not observed
 - data : plasma or insulin concentrations, ...
 - ↪ dynamical models
- Variable informativeness of genetic markers
 - uneven distribution, small sample size of some genotypes
 - ↪ mixed effect models
- Increased size of the genetic data sets toward high throughput screening
 - dimensionality curse $N \ll p$
 - structural correlation along the genome (linkage disequilibrium)
 - ↪ statistical genetics

Genetic association analyses in PGx

- Stepwise procedure
 - commonly used for covariate model building
 - Lehr et al. (2010) adaptation for high throughput screening
- Integrated approach based on penalized regression
 - penalized regression established in animal and plant genetics
 - Lasso (Tibshirani. 1996)
 - HLasso (Hoggart et al. 2008)
 - developed for genome-wise association studies
 - higher effect size once included in the model
- Bayesian variable selection (O'Hara & Sillanpaa 2009)

Stepwise procedure (Lehr et al. 2010)

$$\mathbf{y} = \begin{bmatrix} y_{11} \\ y_{ij} \\ y_{Nn_N} \end{bmatrix}$$

$$SNP = \begin{bmatrix} SNP_{11} & SNP_{1s} & SNP_{1N_s} \\ SNP_{i1} & SNP_{is} & SNP_{iN_s} \\ SNP_{N1} & SNP_{Ns} & SNP_{NN_s} \end{bmatrix}$$

Stepwise procedure (Lehr et al. 2010)

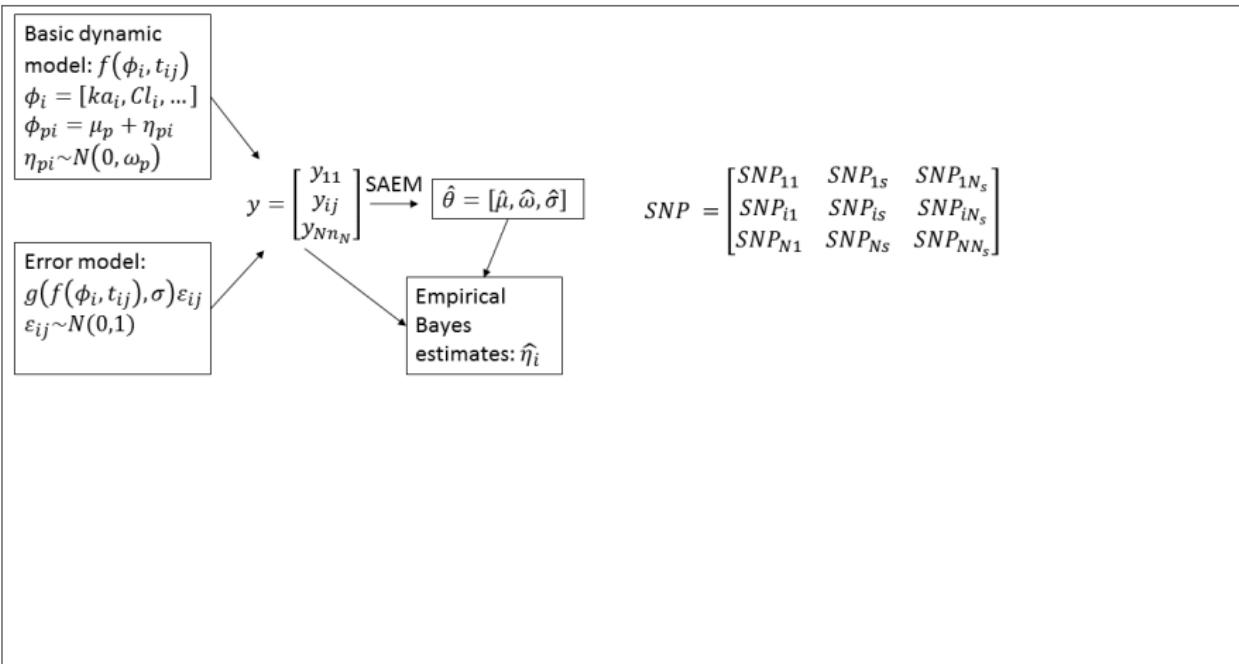
Basic dynamic model: $f(\phi_i, t_{ij})$
 $\phi_i = [ka_i, Cl_i, \dots]$
 $\phi_{pi} = \mu_p + \eta_{pi}$
 $\eta_{pi} \sim N(0, \omega_p)$

$$y = \begin{bmatrix} y_{11} \\ y_{ij} \\ y_{Nn_N} \end{bmatrix} \xrightarrow{\text{SAEM}} \hat{\theta} = [\hat{\mu}, \hat{\omega}, \hat{\sigma}]$$

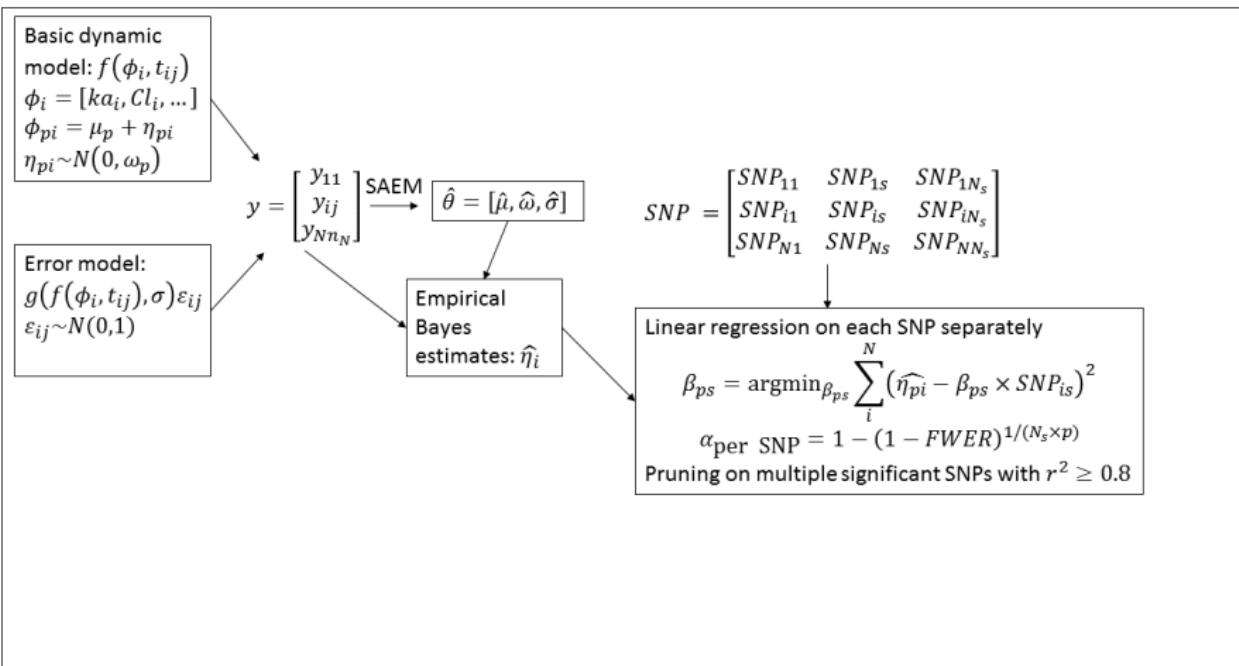
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Error model:
 $g(f(\phi_i, t_{ij}), \sigma) \varepsilon_{ij}$
 $\varepsilon_{ij} \sim N(0, 1)$

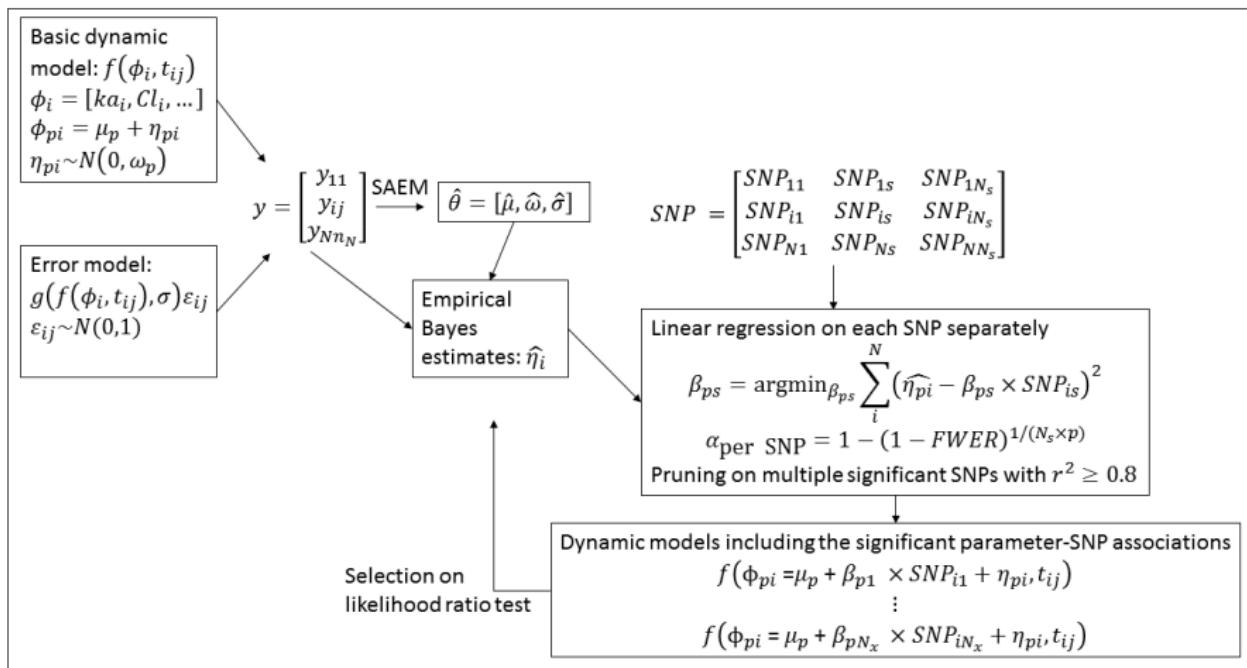
Stepwise procedure (Lehr et al. 2010)



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Stepwise procedure (Lehr et al. 2010)



- SNP selection after estimation of model parameters
- SNP considered independently

Integrated appr. (Bertrand et al. 2015)

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$$SNP = \begin{bmatrix} SNP_{11} & SNP_{1s} & SNP_{1N_s} \\ SNP_{i1} & SNP_{is} & SNP_{iN_s} \\ SNP_{N1} & SNP_{Ns} & SNP_{NN_s} \end{bmatrix}$$

Integrated appr. (Bertrand et al. 2015)

Dynamic model including parameter-SNP associations:
 $f(\phi_i = \mu + \beta \times \text{SNP}_i + \eta_i, t_{ij})$
 $\eta_i \sim N(0, \Omega)$

Error model:
 $g(f(\phi_i, t_{ij}), \sigma) \varepsilon_{ij}$
 $\varepsilon_{ij} \sim N(0, 1)$

$$y = \begin{bmatrix} y_{11} \\ y_{ij} \\ y_{Nn_N} \end{bmatrix} \quad SNP = \begin{bmatrix} SNP_{11} & SNP_{1s} & SNP_{1N_s} \\ SNP_{i1} & SNP_{is} & SNP_{iN_s} \\ SNP_{N1} & SNP_{Ns} & SNP_{NN_s} \end{bmatrix}$$

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SAEM

Stochastic Approximation Expectation-step at iteration k

ϕ'_{ik} drawn from $p(\cdot | y; \theta_k)$, Metropolis Hasting algorithm

$s_{ik} = s_{ik} + \tau_k (\phi'_{ik} - s_{ik-1})$, τ_k sequence of decreasing positive numbers

Maximization-step of μ and β at iteration k

$$(\widehat{\mu}, \widehat{\beta}) = \operatorname{argmin}_{\mu, \beta} \sum_{i=1}^N (s_{ik} - \mu - \beta \times \text{SNP}_i) \Omega^{-1} (s_{ik} - \mu - \beta \times \text{SNP}_i)$$

$$\hat{\theta} = [\hat{\mu}, \hat{\beta}, \hat{\omega}, \hat{\sigma}]$$

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SAEM modified

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Maximization-step of μ and β at iteration k

$(\hat{\mu}, \hat{\beta}) = \operatorname{argmin}_{\mu, \beta} \sum_{i=1}^N (s_{ik} - \mu - \beta \times SNP_i) \Omega^{-1} (s_{ik} - \mu - \beta \times SNP_i) + P(\beta)$

Lasso: $P_\xi(\beta)$ approx. a double exponential prior on β

Hlasso: $P_{\gamma, \lambda}(\beta)$ approx. a normal exponential gamma prior on β

ξ and γ set using an asymptotic approximation

$$\hat{\theta} = [\hat{\mu}, \hat{\beta}, \hat{\omega}, \hat{\sigma}]$$

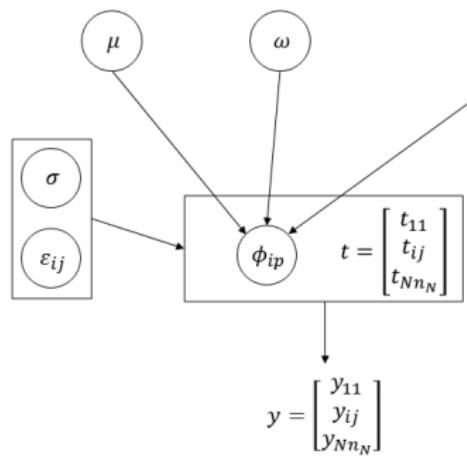
- ↪ Simultaneous SNP selection and estimation of model parameters
- ↪ All parameter-SNP associations considered simultaneously

Bayesian approach (Kuo & Mallick 1998)

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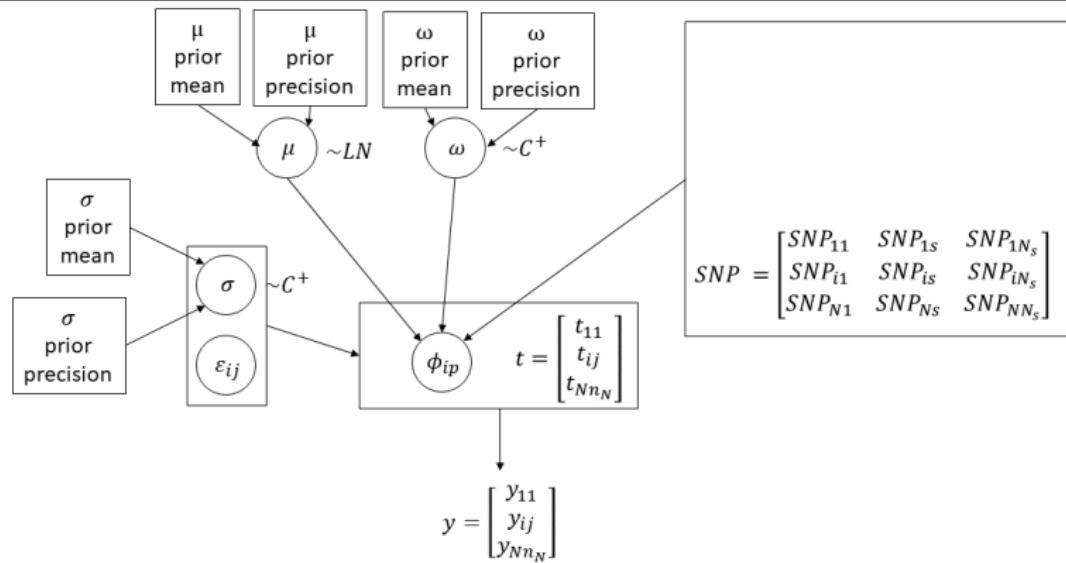
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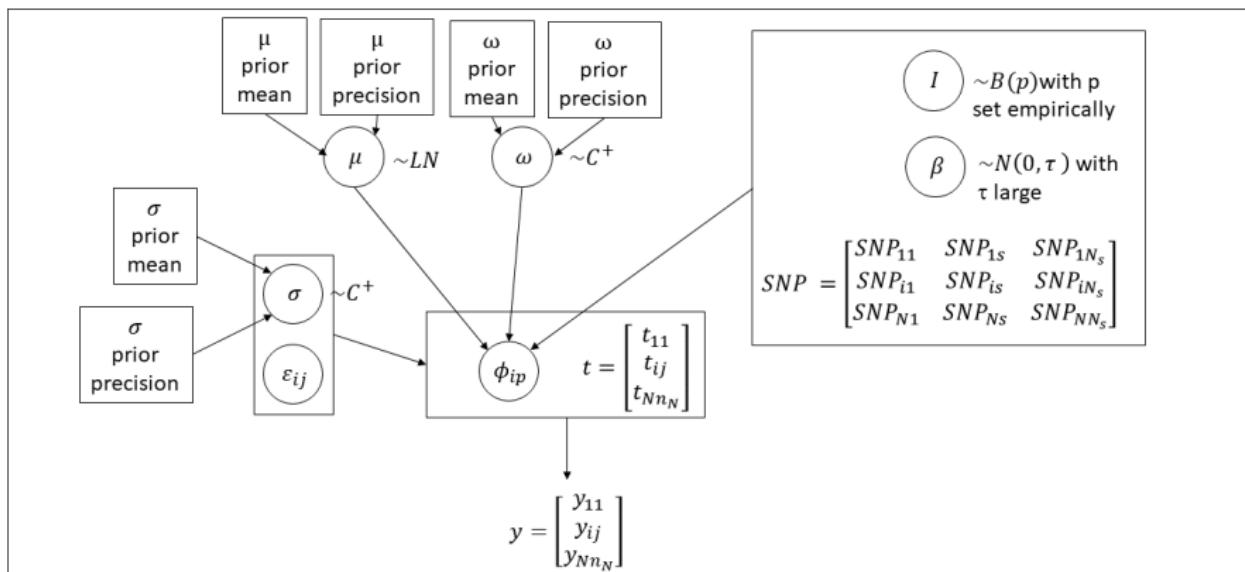


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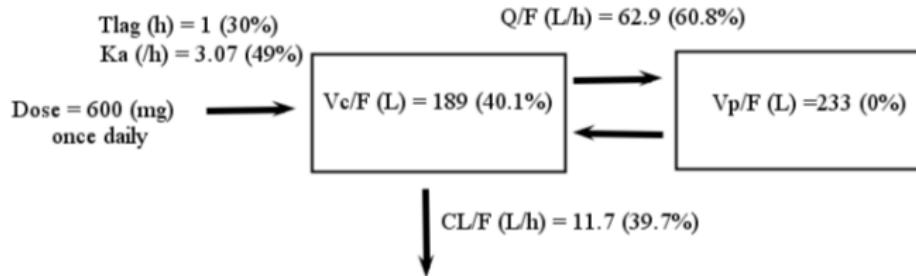
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Objectives

- To evaluate, through a realistic simulation study, the performance to detect a pharmacogenetic association
 - stepwise procedure
 - integrated approach with penalized regression
 - bayesian approach

Genetic and PK data

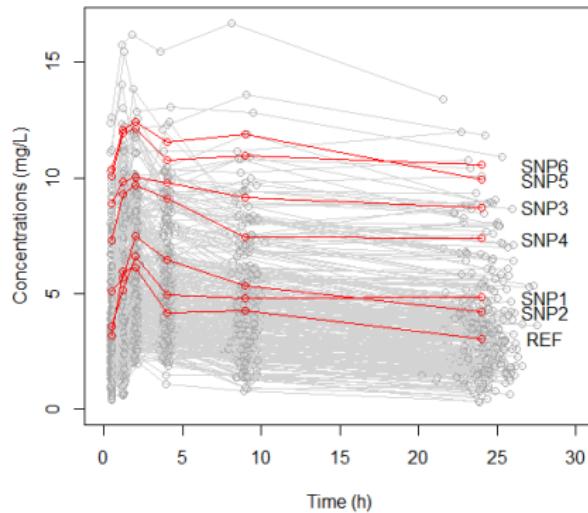
- Generation of genotypes using HAPGEN (Su et al. 2011)
 - $N_s=1227$ snps on 171 genes from the DMET Chip (Daly et al. 2007)
 - 6 [1-56] snps per gene
 - HAPMAP caucasian reference haplotypes
- Pharmacokinetic profiles inspired from real study (Kappelhoff et al. 2005)



- phase II study: $N=300/t=0.5, 1.25, 2, 4, 9, 24\text{h}$
- diagonal variance matrix of random effects
- combined residual error model $g = \sigma_{inter} + \sigma_{prop} f(\phi_i, t_{ij})$

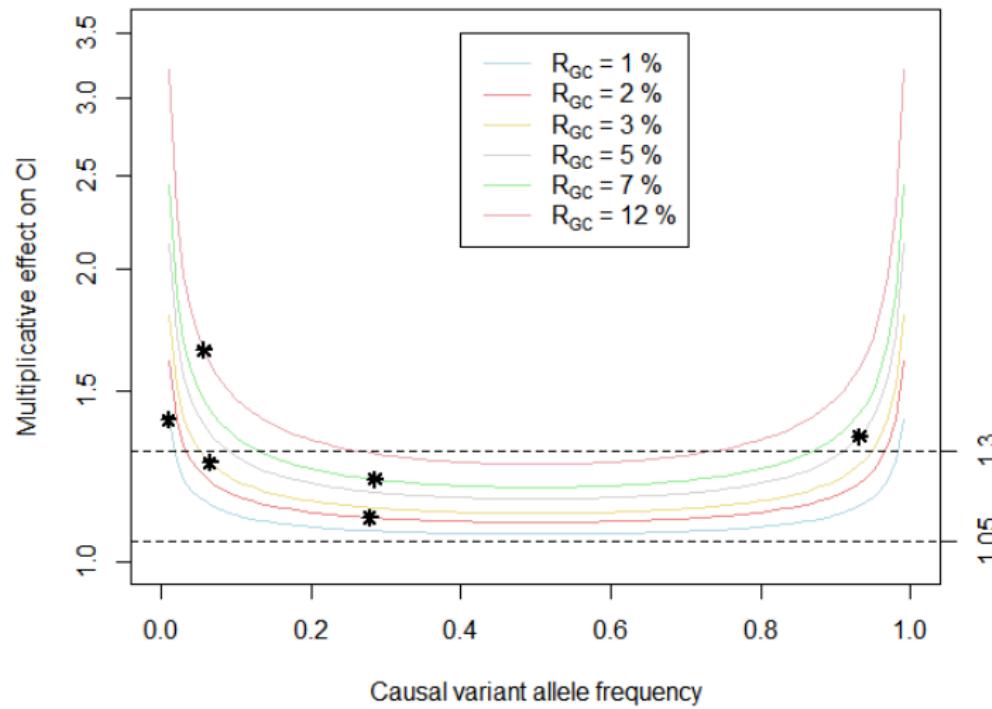
Pharmacogenetic effect - 1/2

- 6 unobserved causal variants explaining 30% of the PK

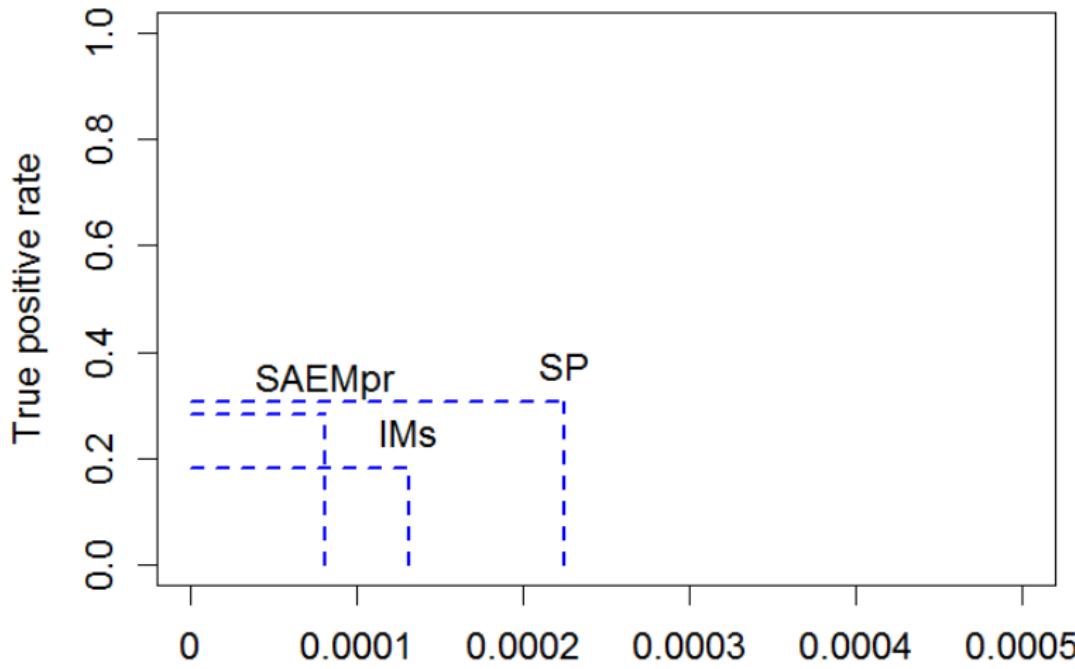


- $\log CI_i = \log CI + \sum_{s=1}^6 \beta_{CI_s} SNP_{s_i} + \eta_{CI_i}$ with $SNP_{s_i} = 0, 1, 2$ i.e. allele dosage model

Pharmacogenetic effect - 2/2



True and false positive rates



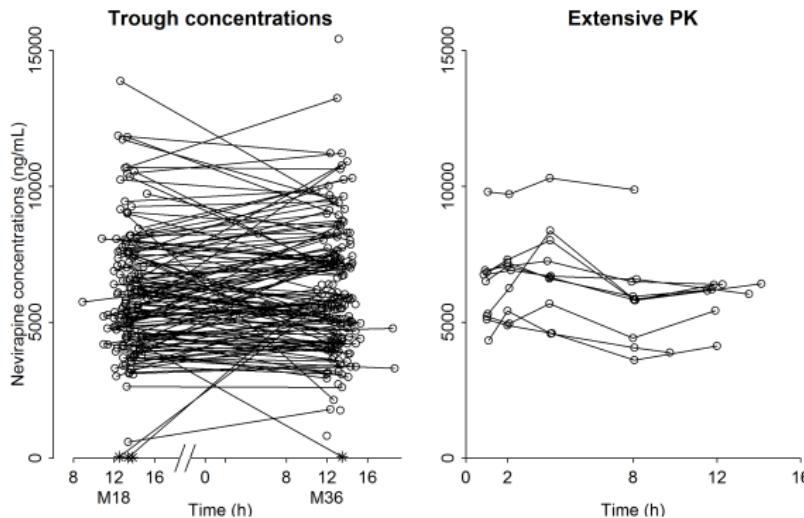
Computing times in hours - mean [range]

Stepwise procedure	0.24 [0.06 - 1.09]
Integrative appr.	1.14[0.83 - 1.61]
Bayesian appr.	19.58 [11.51 - 23.12]

Conclusions

- Feasibility of model-based PGx analyses on a real-case scenario
 - real need of increased sample size compared to classical drug development study designs
- Similar TPR between Integrative appr. and stepwise procedure
 - cost in computing time non-negligible
 - less so on larger data sets (not shown today)
 - better performance when multiple parameter-SNP associations (not shown today)
- Bayesian appr. performance not yet competitive
 - other indicator-based selection and/or shrinkage priors
 - use of Hamiltonian Monte Carlo (rstan package)

PECAN ANRS 12154 (Dr AM Taburet and Prof D Haas)



- 129 patients on up to 3 occasions with 196 markers

Chromosome	3	7	19
Gene	NR1I2 (PXR)	ABCB1 (P-gp)	CYP3A5
Number of markers	49	63	36

- 218 missing polymorphisms with a maximum of 7 per subject

Analyses

- One compartment model with 1st-order absorption and elimination
 - Inter-individual and inter-occasion variability on CL
- 1 Stepwise procedure on empirical Bayes estimates with a priori adjustment for rs3745274
- missing data removed
- 2 Indicator Model selection in JAGS
- missing data imputed from Binomial with empirical allele frequency
- 2 Shrinkage prior in rstan
- missing data imputed beforehand to most common genotype

Stepwise procedure (Bertrand et al. 2012)

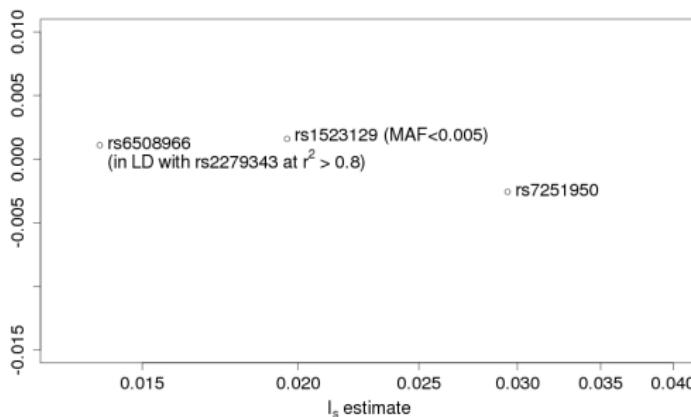
Table 3 Multivariate linear regression for association between genetic variants and nevirapine clearance estimate

Variants	Gene	β	Statistic	P value
<i>n</i> =129 (all participants)				
rs3745274	CYP2B6	-1.606	-6.66	1.01×10^{-9}
rs2687116	CYP3A4	-1.913	-4.68	7.95×10^{-6}
rs2279343	CYP2B6	0.904	4.13	7.07×10^{-5}
rs7251950	CYP2B6	-0.499	-4.23	4.77×10^{-5}
<i>n</i> =128 (without outlier) ^a				
rs3745274	CYP2B6	-1.568	-7.66	6.78×10^{-12}
rs2279343	CYP2B6	0.835	4.50	1.66×10^{-5}
rs7251950	CYP2B6	-0.404	-4.01	1.09×10^{-4}
rs2032582A	ABCB1	0.456	3.60	4.73×10^{-4}

^aData from one individual outlier (extremely high) nevirapine clearance value were censored from this analysis.

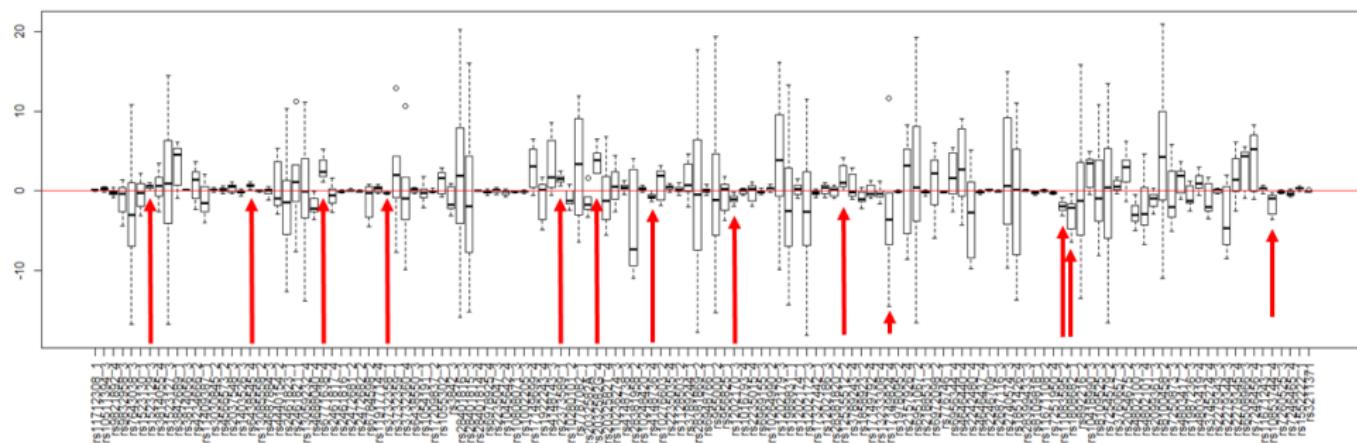
Indicator Model selection in JAGS (Kuo & Mallick

1998)



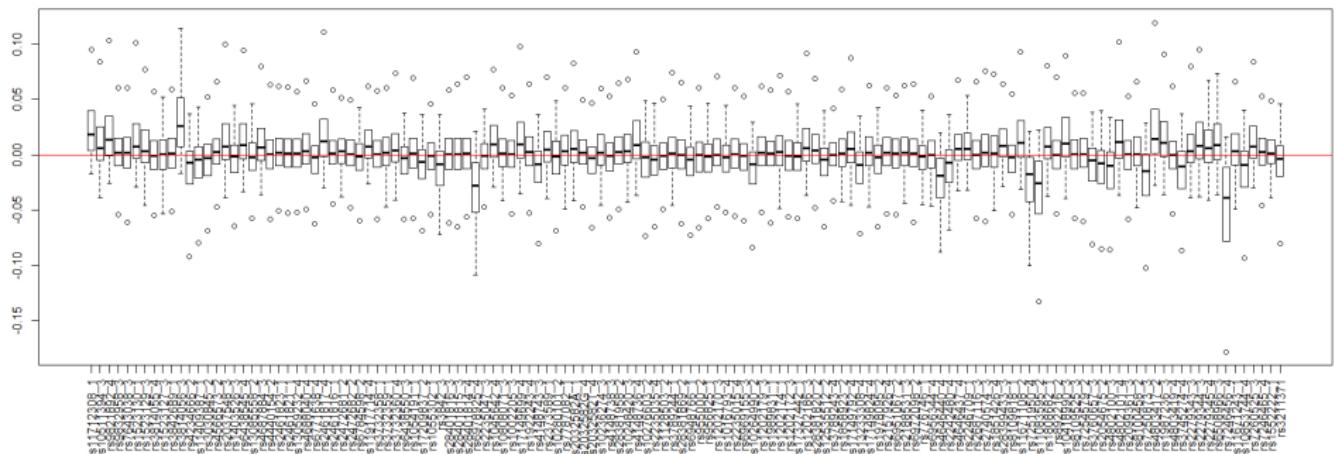
- $\beta = I \times \beta_s$
with $\beta_s \sim N(0, 1)$, $I \sim B(p_{ind})$ and $p_{ind} \sim Beta(.5, .5)$
- 3 chains / nburn=200 / niter=2000 / Neff = [15 -750] /
approx. 5 min

Shrinkage prior in rstan



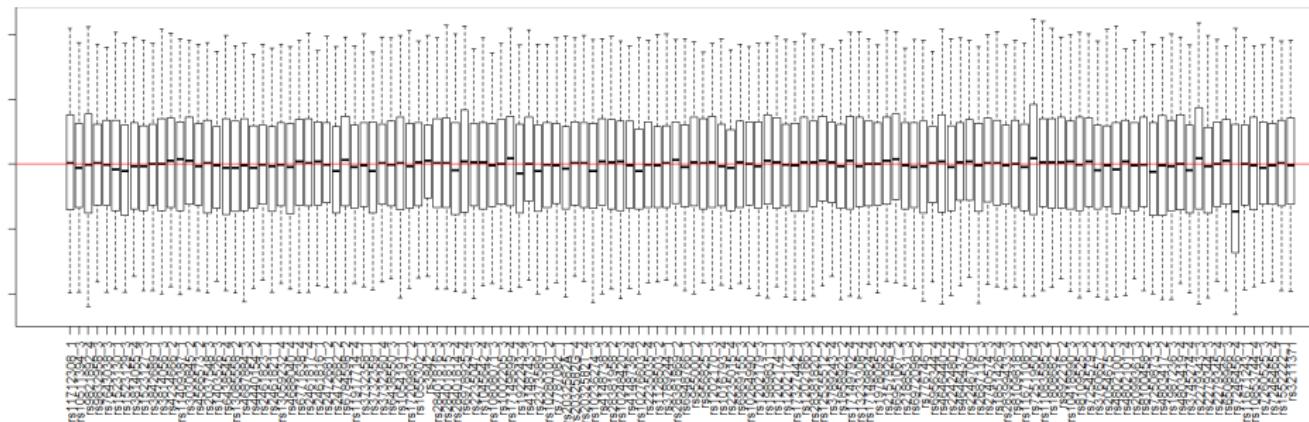
- normal prior (Hoerl & Kennard 1970)
 $\beta \sim N(0, 10)$
- 3 chains / nburn=300 / niter=900 / Neff = [81 - 663] / approx. 20 min

Shrinkage prior in rstan



- Lasso prior (Tibshirani 1996)
 $\beta \sim N(0, \tau)$ with $\tau \sim C^+(0, .01)$
- 3 chains / nburn=300 / niter=900 / Neff = [81 - 663] / approx. 20 min

Shrinkage prior in rstan



- Horseshoe prior (Peltola et al. 2014)
 $\beta \sim N(0, \lambda_i\tau)$ with $\lambda_i \sim C^+(0, 1)$ and $\tau \sim U(0, 1)$
- 3 chains / nburn=300 / niter=900 / Neff = [81 - 663] / approx. 20 min

Conclusion

- Stepwise approach and Bayesian variable selection in JAGs detect similar signals
- Rstan sampling marginally more efficient than JAGs
 - too strong shrinkage of Lasso and Horseshoe ?
- Work in progress with yet a lot to do...
 - run rstan on simulation study
 - investigate how to set prior parameters
 - application to HIV Swiss cohort and EARSII trial on diabetes data

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